

Lecture 6

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1 The Invariance Principle

This lecture will explore a useful generalization of the Central Limit Theorem of probability, which is known as the the Invariance Principle. The Invariance Principle applies in the following setting:

Imagine that we have a function $f : \{-1, 1\}^n \rightarrow \mathbb{R}$. Now imagine that we have i.i.d random variables $x_i \in \{-1, 1\}$ and we compute $f(x_1, \dots, x_n)$. Now $f(x_1, \dots, x_n)$ itself may also be regarded as a random variable. We know that $f : \{-1, 1\}^n \rightarrow \mathbb{R}$ can be expressed as a polynomial in the x_i 's $f = \sum_S \hat{f}_S \prod_{i \in S} x_i$. Given this expression we can ask how much the random variable $f(x_1, \dots, x_n)$ differs from $f(y_1, \dots, y_n)$ where the y_i 's are a different set of independent random variables. For example, the y_i 's could be uniformly distributed on the interval $[-1, 1]$.

It turns out that, just as in the Central Limit Theorem, Gaussian random variables play a central role in the Invariance Principle. We will reserve the notation g_i to refer to a sequence of independent random variables with each g_i distributed according to $\mathcal{N}(0, 1)$ (the Normal distribution with mean 0 and variance 1).

Roughly speaking, we are interested in the question: Does f "notice" the difference between inputs $(x_1, \dots, x_n)$ and $(g_1, \dots, g_n)$. More precisely, we would like to bound the difference between the random variables $f(x_1, \dots, x_n)$ and $f(g_1, \dots, g_n)$. The canonical example of such a result is the Central Limit Theorem

Theorem 1. *The Lindeberg-Levy Central Limit Theorem*

Suppose $\{X_1, X_2, \dots\}$ is a sequence of i.i.d. random variables with $\mathbb{E}[X_i] = \mu$ and $\text{Var}[X_i] = \sigma^2 < \infty$. Then as $n \rightarrow \infty$, the random variables $\sqrt{n} \left(\frac{X_1 + \dots + X_n}{n} - \mu \right)$ converge in distribution to a normal distribution, $\mathcal{N}(0, \sigma^2)$.

In particular, if we define $f(z_1, \dots, z_n) \equiv \frac{1}{n}(z_1, \dots, z_n)$. Then, since $\text{Var}[x_i] = \text{Var}[g_i] = 1$, Theorem 1 implies that $\sqrt{n}f(x_1, \dots, x_n)$, and $\sqrt{n}f(g_1, \dots, g_n)$ both converge in distribution to $\mathcal{N}(0, 1)$ as $n \rightarrow \infty$. Roughly speaking, this means that $\forall t$

$$\lim_{n \rightarrow \infty} \text{Prob}(\sqrt{n}f(x_1, \dots, x_n) \leq t) = \lim_{n \rightarrow \infty} \text{Prob}(\sqrt{n}f(g_1, \dots, g_n) \leq t)$$

In this sense the function $f(z_1, \dots, z_n) \equiv \frac{1}{n}(z_1, \dots, z_n)$ does not "notice" the difference between inputs $(x_1, \dots, x_n)$ and $(g_1, \dots, g_n)$, at least for large n .

However, for arbitrary functions f , it is clear that we will need to be more careful about what question we ask. For example, note that with the simple function $f(z_1, \dots, z_n) \equiv z_1$ we have that $f(x_1, \dots, x_n)$ has a Bernoulli distribution, and $f(g_1, \dots, g_n)$ has a Gaussian distribution, so any reasonable norm will distinguish these two inputs. We revise our question as follows.

Revised Question: Given f that has small influences, that is:

$$\max_{i \in [n]} \text{Inf}_i(f) \leq \tau$$

then does f notice the difference between $x_1, \dots, x_n$ and $g_1, \dots, g_n$?

In this case we have the following theorem:

Theorem 2. *The Mossel-O'Donnell-Oleszkiewicz Invariance Principle*

Let f be a degree d polynomial over $\mathbb{R}$, $f(z_1, \dots, z_n) = \sum_{S \subseteq [n]} \hat{f}_S \prod_{i \in S} z_i$ such that $\sum_S \hat{f}_S^2 = 1$. Assume that $\forall i \text{ Inf}_i(f) \equiv \sum_{S: i \in S} \hat{f}_S^2 \leq \tau$. Then, for any smooth function $\phi: \mathbb{R} \rightarrow \mathbb{R}^+$ with bounded fourth derivative ($|\phi^{(4)}| \leq B$)

$$|\mathbb{E}_{x \in \{\pm 1\}^n} [\phi(f(x))] - \mathbb{E}_{g \in \mathcal{N}(0,1)^n} [\phi(f(g))]| < \text{small}(\tau, \frac{1}{d})$$

Later in the lecture we will discuss a proof of Theorem 2 where $\text{small}(\tau, \frac{1}{d}) = \tau d$. Before we discuss the proof, let us consider an application.

2 Dictatorship vs. τ -Quasirandom Tests

Here a τ -Quasirandom function is a boolean function with all influences smaller than τ just as in 2.

The boolean hypercube is used as a gadget in inapproximability reductions, where proving a completeness, soundness gap of (c,s) for the approximation boils down to devising a (c,s)-Dictatorship vs. Quasirandom test (for boolean functions) which has the following properties:

Completeness: If f is a dictatorship then $\Pr[\text{Test passes}] \geq c - o(1)$

Soundness: If f is quasirandom then $\Pr[\text{Test passes}] \leq s + o(1)$.

Let's consider a test which corresponds to inapproximability of "Max-Cut" or Max-2LIN. This test was proposed by Kindler, Khot, Mossel and O'Donnell, "KKMO".

KKMO "max-cut" test:

Fix $0 < \rho < 1$ (we will also use a parameter δ defined by $\rho = 1 - 2\delta$).

1) Let $x \in \{\pm 1\}^n$ be chosen uniformly at random.

2) Choose $y \in \{\pm 1\}^n$ from the distribution $\mathcal{N}_\rho(x)$ (the distribution which is ρ -correlated with x). That is choose y such that for the i^{th} bit of y (for all i), we have $y_i = x_i$ with probability ρ , and with probability $1 - \rho$ y_i is uniformly random.

3) Accept if $f(x) = f(y)$.

KKMO Completeness: If f is a dictator function (on the i^{th} coordinate say), then

$$\text{Prob}[\text{Test passes}] = \text{Prob}[x_i = y_i] = 1 - \delta$$

KKMO Soundness: If f is quasi-random (meaning all influences are small), then

$$\text{Prob}[\text{Test passes}] = \text{Prob}[f(x) = f(y)] \equiv \text{NS}_\rho(f) \leq \text{NS}_\rho(\text{MAJ}) = \frac{\arccos(\rho)}{\pi}$$

Here we define $\text{NS}_\rho(f) \equiv \text{Prob}[f(x) = f(y)]$. We will now show how the Invariance Principle can be used to calculate $\text{NS}_\rho(\text{MAJ})$ and obtain the final equality above.

A Calculation

Note that

$$\text{MAJ}(x_1, \dots, x_n) = \text{sgn}\left(\sum_{i \in [n]} x_i\right) \approx \text{sgn}(g)$$

and

$$\text{MAJ}(y_1, \dots, y_n) = \text{sgn}\left(\sum_{i \in [n]} y_i\right) \approx \text{sgn}(g')$$

Where g and g' are defined to be normal distributions distributed "like" $\frac{1}{n} \sum_{i \in [n]} x_i$, and $\frac{1}{n} \sum_{i \in [n]} y_i$ respectively. In fact we can use $g \sim \mathcal{N}(0, 1)$, and $g' \equiv \rho g + \sqrt{1 - \rho^2} g''$ where $g'' \sim \mathcal{N}(0, 1)$ and g and g'' are independent. The Invariance Principle tells us that this g and g' will be very close to having the desired distributions.

With this definition of g and g' , straightforward calculus gives

$$\text{NS}_\rho(\text{MAJ}) \approx \text{Prob}[\text{sgn}(g) = \text{sgn}(g')] = \frac{\arccos(\rho)}{\pi}$$

3 Proof of the Invariance Principle

We will now prove the Invariance Principle modulo a certain lemma (Lemma 3 below). We will also discuss the key idea for proving that lemma and use it to obtain a similar (though weaker) result.

Given f as in the statement of Theorem 2 we define

$$f_i \equiv f(g_1, \dots, g_i, x_{i+1}, \dots, x_n)$$

Note that $f_0 = f(x_1, \dots, x_n)$, and $f_n = f(g_1, \dots, g_n)$. Imagine that we have a $\phi : \mathbb{R} \rightarrow \mathbb{R}^+$ as in Theorem 2.

Lemma 3.

$$\forall i, \quad |\mathbb{E}[\phi(f_{i-1})] - \mathbb{E}[\phi(f_i)]| < (Inf_i)^2$$

Using Lemma 3 and the triangle inequality we can now prove Theorem 2.

Proof. Proof of Theorem 2

$$\begin{aligned} |\mathbb{E}[\phi(f(x_1, \dots, x_n))] - \mathbb{E}[\phi(f(g_1, \dots, g_n))]| &= |\mathbb{E}[\phi(f_0)] - \mathbb{E}[\phi(f_n)]| = \left| \sum_{i=0}^{n-1} (\mathbb{E}[\phi(f_i)] - \mathbb{E}[\phi(f_{i+1})]) \right| \\ &\leq \sum_{i=0}^{n-1} |\mathbb{E}[\phi(f_i)] - \mathbb{E}[\phi(f_{i+1})]| \leq \sum_{i=1}^n (\text{Inf}_i)^2 \leq \max_i \text{Inf}_i \left(\sum_{i=1}^n \text{Inf}_i \right) \leq \tau d \end{aligned}$$

Here the first inequality follows by the triangle inequality, the second by Lemma 3, the third inequality is straight forward, and the fourth follows from the fact that f is τ -Quasirandom and degree d by assumption ($\sum_{i=1}^n \text{Inf}_i \leq d$ for degree- d f).

□

Clearly Lemma 3 plays a key role in the proof of the Invariance Principle. We will now give a proof of a weaker version of Lemma 3 which nonetheless contains the key idea of the full proof.

Lemma 4.

$$\forall i, \quad |\mathbb{E}[\phi(f_{i-1})] - \mathbb{E}[\phi(f_i)]| < O(B \cdot 9^d)(\text{Inf}_i)^2$$

Here B is the bound on the fourth derivative of ϕ , and d is the degree of f , just as in Theorem 2.

Proof. Recall Taylor's formula

$$\phi(R + \epsilon) = \phi(R) + \epsilon \phi'(R) + \frac{\epsilon^2}{2!} \phi^{(2)}(R) + \frac{\epsilon^3}{3!} \phi^{(3)}(R) + \frac{\epsilon^4}{4!} \phi^{(4)}(\eta)$$

where η is some point in $\mathbb{R}$.

Now, for a given i , we divide $f(z_1, \dots, z_n)$ up as a polynomial of z_i , so $f(z_1, \dots, z_n) = R(z_1, \dots, z_{i-1}, z_{i+1}, \dots, z_n) + z_i S(z_1, \dots, z_{i-1}, z_{i+1}, \dots, z_n)$. Recall that

$$f_i \equiv f(g_1, \dots, g_i, x_{i+1}, \dots, x_n) = R(g_1, \dots, g_{i-1}, x_{i+1}, \dots, x_n) + g_i S(g_1, \dots, g_{i-1}, x_{i+1}, \dots, x_n)$$

$$f_{i-1} \equiv f(g_1, \dots, g_i, x_{i+1}, \dots, x_n) = R(g_1, \dots, g_{i-1}, x_{i+1}, \dots, x_n) + x_i S(g_1, \dots, g_{i-1}, x_{i+1}, \dots, x_n)$$

So defining random variables $R \equiv R(g_1, \dots, g_{i-1}, x_{i+1}, \dots, x_n)$, and $S \equiv S(g_1, \dots, g_{i-1}, x_{i+1}, \dots, x_n)$ we have that

$$\mathbb{E}[\phi(f_{i-1})] = \mathbb{E}[\phi(R + x_i S)] = \mathbb{E}[\phi(R)] + \mathbb{E}[x_i S \phi'(R)] + \mathbb{E}\left[\frac{(x_i S)^2}{2!} \phi^{(2)}(R)\right]$$

$$\begin{aligned}
& + \mathbb{E} \left[\frac{(x_i S)^3}{3!} \phi^{(3)}(R) \right] + \mathbb{E} \left[\frac{(x_i S)^4}{4!} \phi^{(4)}(\eta) \right] = \mathbb{E} [\phi(R)] + \mathbb{E}[x_i] \mathbb{E} [S \phi'(R)] \\
& + \mathbb{E}[x_i^2] \mathbb{E} \left[\frac{S^2}{2!} \phi^{(2)}(R) \right] + \mathbb{E}[x_i^3] \mathbb{E} \left[\frac{S^3}{3!} \phi^{(3)}(R) \right] + \mathbb{E}[x_i^4] \mathbb{E} \left[\frac{S^4}{4!} \phi^{(4)}(\eta) \right]
\end{aligned}$$

Where the second equality follows from the independence of x_i and the other variables. Very similarly,

$$\begin{aligned}
\mathbb{E}[\phi(f_i)] &= \mathbb{E}[\phi(R + g_i S)] = \mathbb{E} [\phi(R)] + \mathbb{E} [g_i S \phi'(R)] + \mathbb{E} \left[\frac{(g_i S)^2}{2!} \phi^{(2)}(R) \right] \\
& + \mathbb{E} \left[\frac{(g_i S)^3}{3!} \phi^{(3)}(R) \right] + \mathbb{E} \left[\frac{(g_i S)^4}{4!} \phi^{(4)}(\eta') \right] = \mathbb{E} [\phi(R)] + \mathbb{E}[g_i] \mathbb{E} [S \phi'(R)] \\
& + \mathbb{E}[g_i^2] \mathbb{E} \left[\frac{S^2}{2!} \phi^{(2)}(R) \right] + \mathbb{E}[g_i^3] \mathbb{E} \left[\frac{S^3}{3!} \phi^{(3)}(R) \right] + \mathbb{E}[g_i^4] \mathbb{E} \left[\frac{S^4}{4!} \phi^{(4)}(\eta') \right]
\end{aligned}$$

Again the second equality follows from the independence of g_i and the other variables. Now we note that $\mathbb{E}[g_i] = \mathbb{E}[x_i] = 0$, $\mathbb{E}[g_i^2] = \mathbb{E}[x_i^2] = 1$, and $\mathbb{E}[g_i^3] = \mathbb{E}[x_i^3] = 0$. The first two equalities follow by definition of g_i . The third follows because the function $y \rightarrow y^3$ is odd, and both x_i and g_i are have distributions which are symmetric about the origin (they are negative exactly as often as they are positive...). It follows that

$$\begin{aligned}
|\mathbb{E}[\phi(f_{i-1})] - \mathbb{E}[\phi(f_i)]| &= \left| \mathbb{E}[x_i^4] \mathbb{E} \left[\frac{S^4}{4!} \phi^{(4)}(\eta) \right] - \mathbb{E}[g_i^4] \mathbb{E} \left[\frac{S^4}{4!} \phi^{(4)}(\eta') \right] \right| \\
&\leq \left| \mathbb{E}[x_i^4] \mathbb{E} \left[\frac{S^4}{4!} \phi^{(4)}(\eta) \right] \right| + \left| \mathbb{E}[g_i^4] \mathbb{E} \left[\frac{S^4}{4!} \phi^{(4)}(\eta') \right] \right| \leq O(B) \left| \mathbb{E} \left[\frac{S^4}{4!} \right] \right| = O(B) \mathbb{E} [S^4]
\end{aligned}$$

Where the last inequality uses the fact that $|\phi^{(4)}(\eta')|, |\phi^{(4)}(\eta)| \leq B$ by assumption, and the fact that $\mathbb{E}[x_i^4]$ and $\mathbb{E}[g_i^4]$ are both finite constant (not hard to calculate).

For the final step we will apply Hypercontractivity. Note that $S = \frac{d}{dx_i} f_{i-1}$, so that $\mathbb{E}[S^2] = \text{Inf}_i(f_{i-1}) = \sum_{S:i \in S} \hat{f}_S^2$. Clearly, the degree of S is at most d , so by Hypercontractivity we get

$$\mathbb{E} [S^4] \leq (\mathbb{E} [S^2])^2 9^d$$

Using the previous work gives

$$|\mathbb{E}[\phi(f_{i-1})] - \mathbb{E}[\phi(f_i)]| \leq O(B) \mathbb{E} [S^4] \leq O(B) (\mathbb{E} [S^2])^2 9^d \leq O(B \cdot 9^d) (\text{Inf}_i)^2$$

This is the desired result. □

In fact, with a more careful use of Hypercontractivity, and the noisy hypercube function T_ρ the proof of Lemma 4 can be tightened to obtain Lemma 3.