The Chinese University of Hong Kong, Dept. of Electronic Engineering

ELEG 5040: Homework #2

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PROBLEM 1

1. Proof: According to the definition of marginal distribution, we have

$$P(\mathbf{x}) = \sum_{\mathbf{h}} P(\mathbf{x}, \mathbf{h}) = \sum_{\mathbf{h}} \frac{e^{-E(\mathbf{x}, \mathbf{h})}}{Z}$$
$$= \frac{\sum_{\mathbf{h}} e^{-E(\mathbf{x}, \mathbf{h})}}{Z} = \frac{e^{\log \sum_{\mathbf{h}} e^{-E(\mathbf{x}, \mathbf{h})}}}{Z}$$
$$= \frac{e^{-(-\log \sum_{\mathbf{h}} e^{-E(\mathbf{x}, \mathbf{h})})}}{Z}$$
$$= \frac{e^{-\mathscr{F}(\mathbf{x})}}{Z}$$

where $\mathscr{F}(\mathbf{x}) = -\log \sum_{\mathbf{h}} e^{-E(\mathbf{x},\mathbf{h})}$.

Because $\sum_{\mathbf{x}} P(\mathbf{x}) = 1$, we have

$$1 = \sum_{\mathbf{x}} P(\mathbf{x}) = \sum_{\mathbf{x}} \frac{e^{-\mathscr{F}(\mathbf{x})}}{Z} = \frac{\sum_{\mathbf{x}} e^{-\mathscr{F}(\mathbf{x})}}{Z}$$

Thus we get

$$Z = \sum_{\mathbf{x}} e^{-\mathscr{F}(\mathbf{x})}$$

2. Proof: First we derive the expression of negative log-likelihood,

$$-\log P(\mathbf{x}) = -\log \frac{e^{-\mathscr{F}(\mathbf{x})}}{Z}$$
$$= \log Z + \mathscr{F}(\mathbf{x})$$

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Then, we derive the partial grdient wrt θ ,

$$-\frac{\partial \log P(\mathbf{x})}{\partial \theta} = \frac{1}{Z} \frac{\partial Z}{\partial \theta} + \frac{\partial \mathscr{F}(\mathbf{x})}{\partial \theta}$$
$$= \frac{1}{Z} \frac{\partial \sum_{\tilde{\mathbf{x}}} e^{-\mathscr{F}(\tilde{\mathbf{x}})}}{\partial \theta} + \frac{\partial \mathscr{F}(\mathbf{x})}{\partial \theta}$$
$$= -\sum_{\tilde{\mathbf{x}}} \frac{1}{Z} e^{-\mathscr{F}(\tilde{\mathbf{x}})} \frac{\partial \mathscr{F}(\tilde{\mathbf{x}})}{\partial \theta} + \frac{\partial \mathscr{F}(\mathbf{x})}{\partial \theta}$$
$$= -\sum_{\tilde{\mathbf{x}}} P(\tilde{\mathbf{x}}) \frac{\partial \mathscr{F}(\tilde{\mathbf{x}})}{\partial \theta} + \frac{\partial \mathscr{F}(\mathbf{x})}{\partial \theta}$$

thus completes the proof.

3. To generate \mathcal{N} from the model, we start with a training sample **x** (sampled from training samples), and do $|\mathcal{N}|$ Gibbs sampling steps:

$$\mathbf{x}_{1} \sim \hat{P}(\mathbf{x})$$
$$\mathbf{h}_{1} \sim P(\mathbf{h}|\mathbf{x}_{1})$$
$$\mathbf{x}_{2} \sim P(\mathbf{x}|\mathbf{h}_{1})$$
$$\mathbf{h}_{2} \sim P(\mathbf{h}|\mathbf{x}_{2})$$
$$\mathbf{x}_{3} \sim P(\mathbf{h}|\mathbf{x}_{2})$$
$$\mathbf{h}_{3} \sim P(\mathbf{h}|\mathbf{x}_{3})$$
$$\vdots$$
$$\mathbf{x}_{|\mathcal{M}|} \sim P(\mathbf{x}|\mathbf{h}_{|\mathcal{M}|-1})$$

Each Gibbs sampling step in RBM only consists two substeps (sample \mathbf{h} given current \mathbf{x} , and sample \mathbf{x} given current \mathbf{h}), while in a fully-connected Boltzmann Machine, each step we have to sample every node in \mathbf{x} and \mathbf{h} given all the other nodes. This is because in RBM, there are no interactions within \mathbf{x} or \mathbf{h} , so given \mathbf{h} , we can sample all the nodes of \mathbf{x} without affecting each other, and vice versa. The sampling step reduce to two substeps.

PROBLEM 2

Proof: First, let's define some notations:

Let $X = [x_1, x_2, ..., x_N] \in \mathbb{R}^{n \times N}$ be the *N* samples (each is a *n*-dimension vector), $H = [h_1, h_2, ..., h_N] \in \mathbb{R}^{p \times N}$ (p < n) be *N* outputs of hidden units and $Y = [y_1, y_2, ..., y_N] \in \mathbb{R}^{n \times N}$ be the *N* outputs. In auto encoder without non-linear transform, we have:

$$H = W_1 X + w_1 u^t \tag{0.1}$$

$$Y = W_2 H + w_2 u^t \tag{0.2}$$

where $W_1 \in \mathbb{R}^{p \times n}$ and $W_2 \in \mathbb{R}^{n \times p}$ are the two transformation matrices, $w_1 \in \mathbb{R}^p$ and $w_2 \in \mathbb{R}^n$ are the two bias terms, and $u = [1, 1, ..., 1]^t \in \mathbb{R}^N$. The target function we want to minimize is:

$$J = ||X - Y||^2$$

1. In the simplest case, we assume *X* and *Y* has zero means and the bias terms are also zeros:

$$\bar{x} = \frac{1}{N}Xu = \mathbf{0}$$
$$\bar{y} = \frac{1}{N}Yu = \mathbf{0}$$
$$w_1 = \mathbf{0}$$
$$w_2 = \mathbf{0}$$

Then the problem reduce to:

$$J = ||X - W_2 H||^2$$

= ||X - W_2 W_1 X||^2

where $W_2W_1 \in \mathbf{R}^{n \times n}$ but rank p < n. This is equivalent to a typical PCA problem. Let suppose the SVD of *X* has the form of

$$X = U_n \Sigma_n V_n^t$$

where the columns of $U_n \in \mathbf{R}^{n \times n}$ are the eigenvectors of XX^t corresponding to eigenvalues $\lambda_1 \ge \lambda_2 \ge ... \ge \lambda_n$, $\Sigma_n = diag[\sigma_1, \sigma_2, ..., \sigma_n]$ with $\sigma_i = \sqrt{\lambda_i}$ and columns of $V_n \in \mathbb{R}^{N \times n}$ are the eigenvectors of $X^t X$.

Then in a PCA problem we know the optimal solution is projecting samples into the space spanned by the first *p* eigenvectors:

$$W_2 W_1 X = U_p \Sigma_p V_p^t$$

with $\Sigma_p = diag[\sigma_1, \sigma_2, ..., \sigma_p]$ and U_p are formed by the first *p* eigenvectors in U_n . In general, suppose *T* is an arbitrary non-singular $p \times p$ matrix,

$$W_2 = U_p T^{-1}$$

and

$$H = T\Sigma_p V_p^t.$$

The intuition is that *H* is the *p*-dimension PCA projection of *X* plus an arbitrary transformation *T* in the *p*-dimensional space and W_2 is the inverse transformation of *T* plus the projection back to original space. Therefore, the space of hidden units is the same span of first *p* principal components.

2. In general cases,

$$J = ||X - Y||^{2}$$

= $||X - W_{2}H - w_{2}u^{t}||^{2}$
= $tr[(X - W_{2}H - w_{2}u^{t})(X - W_{2}H - w_{2}u^{t})^{t}]$
= $tr[(X - W_{2}H)(X - W_{2}H)^{t}] - 2w_{2}^{t}(X - W_{2}H)u + u^{t}uw_{2}^{t}w_{2}$

To minimize *J* with respect to w_2 , we have:

$$\frac{\partial J}{\partial w_2} = -2(X - W_2 H)u + 2Nw_2 = 0$$
$$\hat{w}_2 = \frac{1}{N}(X - W_2 H)u$$

Therefore, the original problem becomes:

$$J = ||X - W_2 H - w_2 u^t||^2$$

= $||X - W_2 H - \frac{1}{N}(X - W_2 H) u u^t||^2$
= $||(X - X \frac{u u^t}{N}) - W_2(H - H \frac{u u^t}{N})||^2$
= $||X' - W_2 H'||^2$

with $X' = X - X \frac{uu^t}{N} = X - \bar{x}u^t$, which is each sample subtracting the mean vector. So w_2 will ensure that X' has zero mean. In addition, it is easy to show that $\hat{w}_2 = \frac{1}{N}(X - W_2H)u$ also ensures $\bar{y} = \frac{1}{N}Yu = \bar{x}$ and the problem has the same form as before.

PROBLEM 3

1. Proof: Since in RBM hidden nodes are binary, ie $h_j \in \{0, 1\}$, the proof is trivial

$$f_j = \log \sum_{i_j} e^{h_j \mathbf{W}_j \cdot \mathbf{x}}$$
$$= \log (e^{0 \mathbf{W}_j \cdot \mathbf{x}} + e^{1 \mathbf{W}_j \cdot \mathbf{x}})$$
$$= \log (1 + e^{\mathbf{W}_j \cdot \mathbf{x}})$$

2. Let plot $f_j(\mathbf{x}) = \log(1 + e^{\mathbf{W}_j \cdot \mathbf{x}})$ wrt $\mathbf{W}_j \cdot \mathbf{x}$ in Fig 0.1. The function $f'(x) = \log(1 + e^x)$ is called a **softplus** function. When $\mathbf{W}_j \cdot \mathbf{x}$ is a large positive value, f_j approaches *x* asymptotically, and approaches 0 when $\mathbf{W}_j \cdot \mathbf{x}$ is a large negative value. When $\mathbf{W}_j \cdot \mathbf{x}$ is close to zero, f_j has strong non-linearity.

Since $P(\mathbf{x}) \propto \prod_j e^{f_j(\mathbf{x})}$, in each dimension j of hidden units, f_j indicates whether an attribute appears and $\mathbf{W}_j \cdot \mathbf{x}$ classifies the samples into two parts ($\mathbf{W}_j \cdot \mathbf{x} > 0$ and $\mathbf{W}_j \cdot \mathbf{x} < 0$) in this dimension.



Figure 0.1: Softplus Function

- Since P(**x**) ∝ ∏_j e^{f_j(**x**)}, f_j constrains the distribution of **x** along the projection direction **W**_j.. An input sample **x** satisfies a constraint f_i when **W**_i.**x** > 0, and since the hidden units are independent (rows of **W** are independent), **x** can satisfy multiple constraints. P(**x**) is product of all experts (constraints), so P(**x**) is large only if all the constraint are satisfied.
- 4. In a mixture model (sum of experts), the probability is a weighted sum of mixture components (experts). For a Gaussian mixture model in particular, the constraints associated with an expert $N(\mathbf{x}; \mu_i, \Sigma_i)$ can be regarded as a local region measured in **Mahalanobis distance** $d_i = \sqrt{(x \mu_i)^T \Sigma_i^{-1} (x \mu_i)}$. Although probability within a certain Mahalanobis distance decreases as number of dimensionality increases, for a certain number of dimensionality, we could find a local region that contains the majority of samples. For example, in one dimensional example, the 3-sigma distance contains 99.7% of samples.

Since probabilities are positive, a mixture distribution can have high probability for a sample **x** if any one of experts assign high probability to that sample.

PROBLEM 4

1. $\frac{\partial L}{\partial \mathbf{W}_{hz}}$

First we define some symbols and derive some expressions:

$$\frac{d\tanh(x)}{dx} = 1 - \tanh^2(x) \tag{0.3}$$

$$\delta_{i,j} = \begin{cases} 1, & i = j \\ 0, & i \neq j \end{cases}$$
$$\delta_a = \begin{pmatrix} \delta_{1,a} \\ \delta_{2,a} \\ \vdots \\ \delta_{n,a} \end{pmatrix}$$

$$h_t = \tanh(W_{xh}x_t + W_{hh}h_{t-1} + b_h)$$
(0.4)

$$z_t = \operatorname{softmax}(W_{hz}h_t + b_z) \tag{0.5}$$

$$\frac{\partial L_t}{\partial (z_t)_i} = -\frac{\partial \log(z_t)_{y_t}}{\partial (z_t)_i} = -\frac{1}{(z_t)_{y_t}} \frac{\partial (z_t)_{y_t}}{\partial (z_t)_i} = -\frac{1}{(z_t)_{y_t}} \delta_{y_t,i}$$
(0.6)

$$\frac{\partial(z_t)_i}{\partial(W_{hz})_{lk}} = (z_t)_i \delta_{i,l}(h_t)_k - (z_t)_i (z_t)_l (h_t)_k \tag{0.7}$$

Then,

$$\begin{aligned} \frac{\partial L_t}{\partial (W_{hz})_{lk}} &= \sum_i \frac{\partial L_t}{\partial (z_t)_i} \frac{\partial (z_t)_i}{\partial (W_{hz})_{lk}} \\ &= \sum_i -\frac{1}{(z_t)_{y_t}} \delta_{y_t,i} [(z_t)_i \delta_{il}(h_t)k - (z_t)_i (z_t)_l (h_t)_k] \\ &= -\frac{1}{(z_t)_{y_t}} [(z_t)_{y_t} \delta_{y_t,l}(h_t)_k - (z_t)_{y_t} (z_t)_l (h_t)_k] \\ &= -\delta_{y_t,l} (h_t)_k + (z_t)_l (h_t)_k \end{aligned}$$

Therefore,

$$\frac{\partial L}{\partial (W_{hz})_{lk}} = \sum_{t} -\delta_{y_t, l} (h_t)_k + (z_t)_l (h_t)_k$$

And in matrix form,

$$\frac{\partial L}{\partial W_{hz}} = \sum_{t} (z_t - \delta_{y_t}) h_t^T$$

2. $\frac{\partial L}{\partial \mathbf{W}_{hh}}$

$$\begin{split} \frac{\partial(z_t)_i}{\partial(h_t)_p} &= (z_t)_i (W_{hz})_{ip} - (z_t)_i (W_{hz}^T)_{p.} z_t \\ \frac{\partial L}{\partial(z_t)_i} &= \frac{\partial L_t}{\partial(z_t)_i} = -\frac{1}{(z_t)_{y_t}} \delta_{y_t,i} \\ \frac{\partial(h_t)_r}{\partial(h_{t-1})_p} &= [1 - (h_t)_r^2] (W_{hh})_{rp} \\ \frac{\partial L}{\partial(h_t)_p} &= \sum_r \frac{\partial L}{\partial(h_{t+1})_r} \frac{\partial(h_{t+1})_r}{\partial(h_t)_p} + \sum_i \frac{\partial L}{\partial(z_t)_i} \frac{\partial(z_t)_i}{\partial(h_t)_p} \\ &= \sum_r \frac{\partial L}{\partial(h_{t+1})_r} [1 - (h_t)_r^2] (W_{hh})_{rp} + \sum_i -\frac{1}{(z_t)_{y_t}} \delta_{y_t,i} [(z_t)_i (W_{hz})_{ip} - (z_t)_i (W_{hz}^T)_{p.} z_t] \\ &= [\sum_r \frac{\partial L}{\partial(h_{t+1})_r} [1 - (h_t)_r^2] (W_{hh})_{rp}] + [(W_{hz}^T)_{p.} z_t - (W_{hz})_{y_t,p}] \\ \frac{\partial(h_t)_p}{\partial(W_{hh})_{mn}} &= [1 - (h_t)_p^2] \delta_{m,p} (h_{t-1})_n \end{split}$$

In matrix form,

$$\frac{\partial L}{\partial h_t} = W_{hh}^T (1 - diag(h_t)^2) \frac{\partial L}{\partial h_{t+1}} + W_{hz}^T (z_t - \delta_{y_t})$$
$$\frac{\partial L}{\partial W_{hh}} = \sum_t \frac{\partial L}{\partial h_t} \frac{\partial h_t}{\partial W_{hh}} = \sum_t (1 - diag(h_t)^2) \frac{\partial L}{\partial h_t} h_{t-1}^T$$