

Two-Way ANOVA

- Two factors ^a: X_1 has I levels and X_2 has J levels.
- Data y_{ijl} can be displayed in a two-way table with I rows and J columns. The (i, j) th cell contains n_{ij} obs

$$y_{ij1}, \dots, y_{ijn_{ij}}.$$

- A balanced design has $n_{ij} = m$.

^aA **factor** is a categorical predictor with possible values called **levels**.

Possible Models

- The interaction model (the most general model) $\mathcal{M}_{R \times C}$

$$y_{ijl} = \mu + \alpha_i + \beta_j + \gamma_{ij} + e_{ijl}.$$

- The additive model \mathcal{M}_{R+C}

$$y_{ijl} = \mu + \alpha_i + \beta_j + e_{ijl}.$$

- The row-effect model \mathcal{M}_R $y_{ijl} = \mu + \alpha_i + e_{ijl}.$

- The column-effect model \mathcal{M}_C $y_{ijl} = \mu + \beta_j + e_{ijl}.$

- The intercept-only model \mathcal{M}_0 $y_{ijl} = \mu + e_{ijl}.$

The LS Estimates

- We can write the (element-wise) models in matrix form.
- Yes, some are **over-parameterized**. There are multiple ways to reduce parameters, .e.g., we can remove some columns. But no matter which approach we take, the resulting LS estimates should be the same.

- $\mathcal{M}_{R \times C}$ $\hat{y}_{ijl} = \bar{y}_{ij}$.
- \mathcal{M}_{R+C} $\hat{y}_{ijl} = \bar{y}_{i..} + \bar{y}_{.j.} - \bar{y}_{...}$ ^a
- \mathcal{M}_R $\hat{y}_{ijl} = \bar{y}_{i..}$.
- \mathcal{M}_C $\hat{y}_{ijl} = \bar{y}_{.j.}$.
- \mathcal{M}_0 $\hat{y}_{ijl} = \bar{y}_{...}$.

^aThis is true only for balanced design. For unbalanced design, the LS for \mathcal{M}_{R+C} does not have a closed-form expression.

Which model to pick?

- Recall the partial F -test for comparing two **nested** models:

H_0 : a smaller model with p_0 coefficients

H_a : a large model with p_a coefficients

$$F = \frac{(\text{RSS}_0 - \text{RSS}_a) / (p_a - p_0)}{\hat{\sigma}_a^2}$$
$$\sim F_{p_a - p_0, n - p_a} \text{ under the null.}$$

- Suppose we are comparing **three nested** models:

$$\mathcal{M}_1 \subset \mathcal{M}_2 \subset \mathcal{M}_3 \text{ with } \dim p_1 < p_2 < p_3.$$

- We make our decision through the following path:
 - 1) Compare \mathcal{M}_2 vs \mathcal{M}_3 . If \mathcal{M}_3 is selected, stop;
 - 2) otherwise, compare \mathcal{M}_1 vs \mathcal{M}_2 .
- When calculating the F -stat at step 2, we could use $\hat{\sigma}^2$ from \mathcal{M}_2 , i.e., the F -test $\sim F_{p_2-p_1, n-p_2}$ or we could calculate the two F -stats simultaneously using $\hat{\sigma}^2$ from \mathcal{M}_3 (the largest model), so the F -test for the comparison at step 2 would be $F_{p_2-p_1, n-p_3}$.

- Back to the [two-way ANOVA model](#). Due to the hierarchical structure, we make our decision through the following path:
 - 1) Compare $\mathcal{M}_{R \times C}$ vs \mathcal{M}_{R+C} . If select the interaction model, stop; otherwise, go to the next step.
 - 2) Pick one among the remaining four models,

$$\mathcal{M}_{R+C}, \quad \mathcal{M}_R, \quad \mathcal{M}_C, \quad \mathcal{M}_0,$$

depending on whether the row or column effect is significant.

Step 2 is much easier for the balanced design.

The Balanced Design ($n_{ij} = m > 1$)

	SS(Sum Sq)	df
Interaction	$\text{RSS}(\mathcal{M}_{R+C}) - \text{RSS}(\mathcal{M}_{R \times C})$	$(I - 1)(J - 1)$
Row	$\text{RSS}(\mathcal{M}_0) - \text{RSS}(\mathcal{M}_R)$	$I - 1$
Col	$\text{RSS}(\mathcal{M}_0) - \text{RSS}(\mathcal{M}_C)$	$J - 1$
Err	$\text{RSS}(\mathcal{M}_{R \times C})$	$n - IJ$

Three F-tests: SS/df is the numerator, and the dominator is SS/df , i.e., $\hat{\sigma}^2$ from the interaction model

$$\hat{\sigma}^2 = \frac{\text{RSS}(\mathcal{M}_{R \times C})}{(m - 1)IJ}.$$

Selected Model	Interaction	Row	Col
$\mathcal{M}_{R \times C}$	Sig	—	—
\mathcal{M}_{R+C}	No	Sig	Sig
\mathcal{M}_R	No	Sig	No
\mathcal{M}_C	No	No	Sig
\mathcal{M}_0	No	No	No

The Balanced Design ($n_{ij} = m = 1$)

- Only one observation in each cell, so we **cannot fit the interaction model**.
- $\text{RSS}(\mathcal{M}_{R \times C}) = 0$, i.e., the corresponding error variance is 0.
- Then consider \mathcal{M}_{R+C} , instead of $\mathcal{M}_{R \times C}$, to be the largest model.
All the F -tests are the same except that the interaction model is not a candidate model.

The Unbalanced Design

- Compare $\mathcal{M}_{R \times C}$ vs \mathcal{M}_{R+C} : if the F -test is significant, stop.
- If the interaction is not significant, we need to pick one model from

$$\mathcal{M}_{R+C}, \quad \mathcal{M}_R, \quad \mathcal{M}_C, \quad \mathcal{M}_0.$$

- **The difficulty:** to decide whether the column effect is significant, we can
 - a) compare \mathcal{M}_{R+C} vs \mathcal{M}_R (test the column effect **given that the row effect has been included**), or
 - b) compare \mathcal{M}_C vs \mathcal{M}_0 (test the column effect **given that the row effect is not included**).

- For a balanced design, these two tests turn out to be the same, but that's not the case for the unbalanced design.
- In most cases, it is not difficult to make a consensus decision.
- Let's take a look of the two ANOVA tables.

	SS(Sum Sq)	df
<i>C</i>	$RSS(\mathcal{M}_0) - RSS(\mathcal{M}_C)$	$J - 1$
<i>R C</i>	$RSS(\mathcal{M}_C) - RSS(\mathcal{M}_{R+C})$	$I - 1$
Interaction	$RSS(\mathcal{M}_{R+C}) - RSS(\mathcal{M}_{R \times C})$	$(I - 1)(J - 1)$
Err	$RSS(\mathcal{M}_{R \times C})$	$n - IJ$
TSS	$RSS(\mathcal{M}_0)$	$n - 1$

	SS(Sum Sq)	df
<i>R</i>	$\text{RSS}(\mathcal{M}_0) - \text{RSS}(\mathcal{M}_R)$	$I - 1$
<i>C R</i>	$\text{RSS}(\mathcal{M}_R) - \text{RSS}(\mathcal{M}_{R+C})$	$J - 1$
Interaction	$\text{RSS}(\mathcal{M}_{R+C}) - \text{RSS}(\mathcal{M}_{R \times C})$	$(I - 1)(J - 1)$
Err	$\text{RSS}(\mathcal{M}_{R \times C})$	$n - IJ$
TSS	$\text{RSS}(\mathcal{M}_0)$	$n - 1$

Selected Model	R	$R C$	C	$C R$
\mathcal{M}_0	×	×	×	×
\mathcal{M}_{R+C}	—	✓	—	✓
	✓	—	—	✓
	—	✓	✓	—
\mathcal{M}_R	✓	✓	—	×
	✓	×	×	×
\mathcal{M}_C	—	×	✓	✓
	×	×	×	✓
\mathcal{M}_{R+C}	×	✓	×	×
	×	×	×	✓

	SS(Sum Sq)	df	F -stat
R	20	10	$2 \sim F_{10,50}$
$C \mid R$	16	8	$2 \sim F_{8,50}$
C	11	8	$1.375 \sim F_{8,50}$
$R \mid C$	25	10	$2.5 \sim F_{10,50}$
$R + C$	36	18	$2 \sim F_{18,50}$
Err	50	50	$\hat{\sigma}^2 = 1$