ECE209AS (Winter 2021)

Lecture 2: Making Inferences from Sensor Data (The Pre Deep Learning Era)

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Sensor Measurement \neq State

- Challenge: sensors themselves do not provide environment state
 - E.g. sensors do not say "There is Mr. Smith sitting on a chair and wearing a black suit"
- Rather, sensors may tell
 - Light level
 - ► Color
 - Whether it is touching something in an area
 - Sound level
 - Distance to nearest object
 - ► Etc.
- Sensors measure physical quantities
 - Need to be processed to be useful
- Same physical quantity may be measurable by different sensors • Can help improve accuracy in the presence of error and noise

Physical Property Sensing Technology \rightarrow bump, switch Contact \rightarrow ultrasound, radar, IR Distance \rightarrow Light Level photocells, cameras \rightarrow Sound Level microphones \rightarrow Strain strain guages \rightarrow Rotation encoders, potentiometers \rightarrow Acceleration accelerometers, gyroscopes \rightarrow Magnetism compasses \rightarrow Smell chemical sensors \rightarrow Temperature thermal, IR \rightarrow inclinometers, gyroscopes Inclination \rightarrow Pressure pressure guages \rightarrow Altitude altimeters \rightarrow

Some sensors and the information they measure



Example: Measuring distance to an object

- Ultrasound
 - time of flight
- Infra-red return signal intensity
- Two cameras
 - ▶ stereo
- Single camera
 - using perspective + assumption about environment structure
- Laser + fixed camera
 - triangulate distance
- Laser-based structured lighting (overlay grid pattern) + fixed camera distance from distortion in pattern
- Others?



Example: Detecting people

- Use a camera?
 - Camera/vision is a very powerful modality Intensity, color, texture, shape etc.
 - But very costly in processing
- Other ways: using sensors simpler than vision
 - Temperature: search for temperature ranges corresponding to human body temperature
 - Movement: if everything else is static, movement means people
 - Color: look for colors corresponding to human skin or clothes/uniforms
 - Distance: if an otherwise open distance range becomes blocked, there is likely a human being
- Often simpler sensors are enough
 - ► E.g. burglar alarm
 - can't distinguish humans from other animals, but non-human burglars are rare
 - Plus, can help improve accuracy of vision

Activity Detection: Ambient Sensors in the Environment

- Passive Infrared (PIR)
- Magnetic Door/Window
- Temperature, Light, Humidity
- Vibration
- Pressure
- RFID
- Camera
- Microphone
- Electric meter
- Water meter
- Coverage: near-field vs far-field

















Activity Detection: Wearable Sensors on (or in) Body

	Inertial	25 -	Na
		20 -	
		15 -	
		10 -	
		5 -	
•	PhysiologicalHeart Rate	0 -	20
	Heart Rate Variability		
	Breathing Rate		
	Galvanic Skin Response		
	Blood Pressure		
	► EMG, EEG, etc.		







Example Deployment: Ambient + Wearable Sensors







Sensor Measurements for Human Activities



Cook & Krishnan. Activity Learning: Discovering, Recognizing, and Predicting Human Behavior from Sensor Data. 2015.

Sweeping







Common Sensor Analytics

- Inferring latent states
- Forecasting
- Anomaly Detection
- Missing Value Imputation
- Clustering





Formalizing the Problem

- We need to track and predict the state of a *dynamic* environment in the presence of *uncertainty*
 - system in a building, etc.
 - Unlike static situations, such as analyzing an X-ray image or diagnosing a faulty engine
- Notation
 - $S_t = \text{set of unobservable state variables at time } t$ - e.g. sitting, walking, running, biking, etc.
 - $O_t =$ set of observable evidence variables at time t
 - e.g. measurements from accelerometer, gyroscope, PPG, camera etc.
 - may be multimodal and from spatially distributed sensors
- Simplification of time

 - allows treating time t as an index so that time goes as t = 0, 1, 2, ...
 - notation: $X_{a:b} = X_a, X_{a+1}, \dots, X_{b-1}, X_b$
 - Iater in the course we will relax this regular synchronous sampling

• E.g. managing health of a diabetic individual, accounting for time spent in different activities, control HVAC

• time t is discretized into fixed step-size Δ that is identical for S_t and O_t , and steps aligned for S_t and O_t

Physics vs Data Drive Approaches

- Physics of the environment: equations governing the evolution of the states Physics of the sensor: equations governing mapping of states to observation
- If these are known, we can solve an inverse problem to get state from obersvations
- But:
 - Physics may not be known at all, or may be complex to compute, or highly uncertain Environment and sensor characteristics may be dynamic
- Alternative: data drive "machine learning" approaches

Wearable Sensor-based Activity Recognition

- User performing activities belonging to a predefined set $A = \{A_i\}_{i=1}^m$ where *m* is the number of activity types
- There is a sequence of sensor reading that captures the activity information $\mathbf{o}_{1:n} = [\mathbf{o}_1, \mathbf{o}_2, \cdots \mathbf{o}_t, \cdots \mathbf{o}_n]$ where \mathbf{o}_t is the sensor reading (in general a vector) at time t
- We need to build a model \mathcal{F} to predict the activity sequence based on sensor reading **s**

$$\hat{A}_{1:n} = \mathscr{F}(\mathbf{o}_{1:n}), \quad \forall j \, \hat{A}_j \in A$$

while the true activity sequence (ground truth) is denoted as $A_{1\cdot n}^*, \quad \hat{A}_i^* \in A$

- Goal is to learn \mathcal{F} by minimizing the discrepancy between predicted • activity sequence $\hat{A}_{1:n}$ and the ground truth activity sequence $A^*_{1:n}$
- *n* governs the algorithmic latency in inferring the activity state



Inferring Activity State Using Conventional Patter Recognition

- reading \mathbf{o}_i to a *d*-dimensional feature vector $\Phi(\mathbf{o}_i) \in \mathbb{R}^d$, so that loss function is $\mathscr{L}(\mathscr{F}([\Phi(\mathbf{o}_i)]_{i=1}^n), A_{1\cdot n}^*)$



Goal is to learn \mathscr{F} by minimizing the discrepancy between predicted and the ground truth activity sequence Typically, during training a positive loss function $\mathscr{L}(\mathscr{F}(\mathbf{0}_{1:n}), A^*_{1:n})$ is constructed to reflect this discrepancy • \mathscr{F} does not take $\mathbf{0}_{1:n}$ as an input, but rather assumes there is a projection function Φ that projects each sensor

Common Simplifications

- O_{1:n} grouped into windows of some time duration
 - Activity type assumed constant within each window
 - Activity in a window assumed independent of activities in other window
 - Simplifies \mathcal{F} and makes it easier to deal with irregular and missing observations
 - But also creates problems
 - Activity changes misaligned with segment boundaries
 - Activity state may change multiple times within the segment if k is large
 - Ignores temporal correlations between activities in different segments
- The assumption that all sensors contributing to $\mathbf{0}_t$ have the same sample rate is not correct In fact, sampling rate of a sensor may not even be fixed
- - e.g. in Android the sensor sampling rate is an advice to the OS but it is not guaranteed that timestamps will be equally spaces

 - Preprocessing done to address this before extracting features - it also takes care of missing data, calibration etc.



Features

- Two types of observations from sensors Sensors that generate time series of periodic numeric measurements of state - e.g. periodic sampling of temperature Sensors that generate event reports about change in state - e.g. send event when there is motion
- aspects of the problem
- Quality of the learned model depends on expressiveness of features

Features computed over windows of observations basically get rid of the temporal

Features

Sample of window events from the sweeping activity

Time	ID	Message	12:50:18.119112	HIP	(1072,2052,1559,1880,1796,1949)
12:50:13.102682	Dustpan	MOVED	12:50:18.341054	ARM	(1071,1889,1978,1780,1721,1982
12:50:13.149904		(1019,1921,1520,1850,1800,1898)	12:50:18.452039	M017	ON
2:50:13.367222		(1584,2402,2318,2040,1838,1736)	12:50:18.790212	HandSoap	MOVED
			12:50:19.103787	HIP	(1159,2026,1598,1863,1744,1951
L2:50:14.128257		(973,1951,1545,1839,1918,1900)	12:50:19.349545	ARM	(1043,1330,1688,1676,1825,1872)
L2:50:14.217971	Duster	MOVED	12:50:20.101233	HIP	(986,1966,1573,1838,1727,1921)
12:50:14.357487	ARM	(948,1851,1837,1886,1820,2028)	12:50:20.334268	ARM	(1007,1674,1700,1702,1868,1916
12:50:15.129119	HIP	(1055,1745,1792,1840,1814,1872)	12:50:20.741536	HandSoap	STILL
12:50:15.873883	ARM	(926,1867,2119,1751,1853,1863)	12:50:21.097739		(981,2074,1527,1853,1724,1930)
12:50:16.134036	HIP	(1047,2137,1487,1819,1594,1992)	12:50:21.342635		(1208,2346,1888,2217,1234,1483)
12:50:16.235281	Broom	MOVED	12:50:21.895284		2.8227
12:50:16.36758	ARM	(1055,1677,2182,1705,1613,1862)	12:50:22.090522		(1149,2003,1701,1857,1615,1949)
12:50:17.001771	M018	ON	12:50:22.339887		(1134,2594,2088,1894,1845,1872)
12:50:17.11736	HIP	(997,1970,1625,1752,1438,1907)	12:50:22.412985	HandSoap	
12:50:17.345406	ARM	(1189,1935,2232,1840,1682,1813)	12:50:22.8739	Dustpan	STILL

- Characteristics of sensor event sequence
 - time of occurrence
 - time of day, day of week etc.
 - sequence duration
 - 12:50:22.8739-12:50:13.102682

- Characteristics of sensor event sequence
 - time of occurrence
 - time of day, day of week etc.
 - sequence duration
 - 12:50:22.8739-12:50:13.102682
- Characteristics of discrete sensor values
 - bag of sensors
 - set of events with associated frequencies
 - elapsed time since last event from sensor

Discrete-Based Features								
Sensor Counts		Sensor Elapsed Times						
Dustpan	1	Dustpan	97,712					
Duster	1	Duster	86,559					
Broom	1	Broom	66,386					
Hand soap	2	Hand soap	40,837					
M017 (Kitchen)	2	M017 (Kitchen)	44,219					
M018 (Sink)	1	M018 (Sink)	58,721					
Burner	1	Burner	9,786					
All other sensors	0	All other sensors	?					







- Characteristics of sensor event sequence
 - time of occurrence
 - time of day, day of week etc.
 - sequence duration
 - 12:50:22.8739-12:50:13.102682
- Characteristics of discrete sensor values
 - bag of sensors
 - set of events with associated frequencies
 - elapsed time since last event from sensor
- Statistical features over a time window



Plot of acceleration (Ax,Ay,Az) and rotational velocity (RxRy,Rz) values from the HIP IMU from sweeping activity data sample

- Characteristics of sensor event sequence
 - time of occurrence
 - time of day, day of week etc.
 - sequence duration
 - 12:50:22.8739-12:50:13.102682
- Characteristics of discrete sensor values
 - bag of sensors
 - set of events with associated frequencies
 - elapsed time since last event from sensor
- Statistical features over a time window
- Activity context features
 - previous activity
 - previous dominant sensor
 - weighted features from previous window



Common Statistical Features: Temporal

- Max, Min
- Sum, Mean, Median
- Mean Absolute Deviation, Median Absolute Deviation
- Standard Deviation
- Coefficient of Variation
- Zero Crossings
- Percentiles, Inter-quartile Range
- Square Sum of Percentile Observations
- Histogram
- Skewness (degree of asymmetry of distribution)
- Kurtosis (peakiness of distribution around mean)
- Correlation (across dimensions)
- Autocorrelation
- Signal Energy, Log Energy, Signal Power
- Signal Magnitude Area
- Peak-to-Peak Amplitude
- Time between Peaks

Hip Ax C	Continuous	-Based Feature	Values for	Sweeping Act	ivity
Max	1,159.00	CV	0.09	SqSumPt(80)	8,272,
Min	973.00	ZC	5	BD(1)	0.5
Sum	10,438.00	PT(20)	981.01	BD(2)	0.3
Mean	1,043.00	PT(50)	1019.01	BD(3)	0.2
Median	1,033.00	PT(80)	1072.01	Skewness	0.83
MeanAbsDev	19.12	IQ	91.00	Kurtosis	-0.48
MedAbsDev	25.50	SqSumPt(20)	1,909,090	Corr(A <i>x</i> ,A <i>y</i>)	0.37
StDev	66.98	SqSumPt(50)	4,913,656	AC ₁	-0.20

Ax Signal-Based Features for Sweeping Activity									
Signal energy	10,935,556.0	P2PA	186.0						
Log signal energy	200.5	TBPeaks	3.5						
Power	1,093,555.6	NumPeaks	3						

Common Statistical Features: S

- Spectral Centroid
- Spectral Energy
- Spectral Entropy





$$=\sum_{i=1}^{N}\mathcal{F}(n)^{2}$$

$$T(n) = \frac{\mathcal{F}(n)}{\sum_{i=1}^{N} \mathcal{F}(n)}$$

Spectral Features fr Window	Spectral Features from Hip Ax Data for the Sweeping Activity Window							
of 10 Events								
Spectral centroid	4.9550	Spectral energy	1007864					
Spectral entropy	1.8238	Normalized spectral energy	0.2002					

Common forms of \mathcal{F}

- Naive Bayes
- Linear
- Linear with Non-linear Kernels

sick

yes

sick

X1<0.39

X2<0.40?

ves

 $x_1 \sim w$

- Support Vector Machines
- Decision Tree
- Random Forests
- Bayesian Network
- HMM
- CRF
- Being replaced by neural networks in many cases





What if there are multiple sensors?

- Complementary vs overlapping information
- Three approaches
 - Data fusion
 - average sensor measurements
 - weighted by trust/quality/variance
 - Feature fusion
 - concatenate feature vectors
 - dimensionality reduction methods in case of overlap
 - Classifier fusion
 - simple voting
 - weighted majority voting
 - highest quality (lowest uncertainty)
 - learn a second level classifier
 - Bayesian
 - summing K independent classifiers P(C

$$C|S_1, \dots, S_K) = \frac{P(S_1, \dots, S_K | C) P(C)}{P(S_1, \dots, S_K)} \qquad C^* = \operatorname{argmax}_c \left\{ P(C = c) \prod_{k=1}^K P(S_k = s_k) \right\}$$



Supervised Learning of parameters of model ${\mathcal F}$





A Simple but Very Useful Classifier: Naive Bayes

- Notation
 - X: feature vector of dimension D over a window
 - Y: activity label
- Objective: given X, find Y that maximizes the likelihood P(Y|X)
- Learning $P(X^1, X^2, \dots, X^D | Y)$ from training data would need lots of data
- Assumes that data attributes are conditionally independent given class label $P(X^1, X^2, \dots, X^D | Y) = \prod_{d=1}^D P(X^d | Y)$
- Then:

$$P(Y = k | x^1, x^2, \dots, x^D) = \frac{\prod_{d=1}^{D} P(x^d | Y = k) P(Y = k)}{\sum_j \prod_{d=1}^{D} P(x^d | Y = j) P(Y = j)}$$

$$P(Y|X) = \frac{P(X|Y)P(Y)}{P(X)}$$

$$Y = \operatorname{argmax}_{k} \frac{\prod_{d=1}^{D} P(x^{d} | Y = k) P(Y = k)}{\sum_{j} \prod_{d=1}^{D} P(x^{d} | Y = j) P(Y = j)} = \operatorname{argmax}_{k} \prod_{d=1}^{D} P(x^{d} | Y = k) P(X = k)$$





Naive Bayes Classifier (NBC) contd.

 Learning model parameters from training data using maximum likelihood estimators

 $P(X^{i} = v_{j}|Y = k) = \frac{\text{#data samples with } X^{i} = v_{j} \text{ and } Y = k}{\text{#data samples with } Y = k}$

problem: if no data for a given value of X then likelihood probability 0 solution: smoothing (M is the # of distinct values X can take)

$$P(X^{i} = v_{j}|Y = k) = \frac{(\text{#data samples with } X^{i} = k)}{(\text{#data samples with } X^{i})}$$

 $P(Y = k) = \frac{(\text{#data samples with label } k) + l}{(\text{#data samples}) + lK}$

What if features are continuous instead of categorical?

 $\frac{v_j \text{ and } Y = k) + l}{Y = k) + lM}$



Applying NBC

X	x_1	x_2	x_3	x_4	x_5	x_6	<i>x</i> ₇
Day of the week (X^1)	1	1	1	2	2	2	2
Time of the day in hours (X^2)	2	7	20	6	10	14	21
Bathroom sensor count (X^3)	13	10	5	11	3	12	3
Medicine cabinet sensor count (X^4)	0	2	6	1	4	1	9
Energy in the Hip accelerometer Z axis (X^5)	455	500	200	506	207	521	293
Activity label (Y)	2	1	3	1	4	1	3

- Probability of attribute X¹ (day of week, values in 1-7) taking value 1 given that activity is Personal Hygiene (Y=1): $P(X^1 = 1 | Y = 1) = \frac{1}{3}$
- Prior probability that activity is Personal Hygiene
- Similarly P(X1=1|Y=3) and P(Y=3) are 1/3 and 2/7

$$P(Y=1): P(Y=1) = \frac{3}{7}$$



NBC with Continuous Features

- Typical approach is to model P(Xⁱ|Y=j) as a Gaussian Distribution $N(\mu_i^i, \sigma_i^{2^i})$
- Estimate the parameters of the Gaussian Distribution as: $\mu_j^i = E[X^i | Y = j]$

$$\sigma_{j}^{2^{i}} = E[(X^{i} - \mu_{j}^{i})^{2}|Y =$$

• This estimate can be done from training data using maximum likelihood estimation process:

$$\mu_j^i \approx \frac{\sum_{m:y_m=j} x_m^i}{\text{#data samples with acti}}$$

$$\sigma_j^{2^i} \approx \frac{\sum_{m:y_m=j} (x_m^i - \mu_j)}{\text{#data samples with act}}$$

j]

ivity label j

1.2

tivity label j



Applying NBC contd.

XDay of the week (X^1) Time of the day in hours (X^2) Bathroom sensor count (X^3) Medicine cabinet sensor count (X^3) Energy in the Hip accelerometer (X^5)

Activity label (Y)

• Consider X⁵ (Energy in hip accelerometer Z axis)

$$\mu_1^5 = \frac{500 + 506 + 521}{3} = 509$$
$$\mu_3^5 = 246.5$$

-

Cook & Krishnan. Activity Learning: Discovering, Recognizing, and Predicting Human Behavior from Sensor Data. 2015.

	x_1	x_2	x_3	x_4	x_5	x_6	x_7
	1	1	1	2	2	2	2
	2	7	20	6	10	14	21
	13	10	5	11	3	12	3
(X^4)	0	2	6	1	4	1	9
er Z axis	455	500	200	506	207	521	293
	2	1	3	1	4	1	3

$$\sigma_1^{2^5} = \frac{(500 - 509)^2 + (506 - 509)^2 + (521 - 509)^2}{3}$$

 $\sigma_3^{2^5} = 2162.5$



Applying NBC contd.

X	_		
A	٦	,	
	. A		

Day of the week (X^1)

Time of the day in hours (X^2)

Bathroom sensor count (X^3)

Medicine cabinet sensor count (

Energy in the Hip accelerometer (X^{5})

Activity label (Y)

P(Y = 1).

• Given a test sample:

 $P(Y = 3 | x^1 = 1, x^5 = 250) = P(x^1 = 1 | Y = 3) \times P(x^5 = 250 | Y = 3) \times P(Y = 3)$ $X^{1} = 1$ X⁵=250

Cook & Krishnan. Activity Learning: Discovering, Recognizing, and Predicting Human Behavior from Sensor Data. 2015.

	x_1	x_2	<i>x</i> ₃	x_4	x_5	x_6	<i>x</i> ₇
	1	1	1	2	2	2	2
	2	7	20	6	10	14	21
	13	10	5	11	3	12	3
(X^4)	0	2	6	1	4	1	9
er Z axis	455	500	200	506	207	521	293
	2	1	3	1	4	1	3

$$|x^{1} = 1, x^{5} = 250) = P(x^{1} = 1|Y = 1) \times P(x^{5} = 250|Y = 1) \times P(Y = 1)$$

$$= \frac{1}{3} \times (8.04 \times 10^{-189}) \times \frac{3}{7} = 1.419 \times 10^{-189}$$

$$=\frac{1}{3} \times 0.086 \times \frac{2}{7} = 8.1905 \times 10^{-4}$$

Since $P(Y = 3|x^1 = 1, x^5 = 250) > P(Y = 1|x^1 = 1, x^5 = 250)$, the NBC will label this data point as Take Medicine (Y = 3).

NBC Summary

- Probabilistic approach to classification that is popular for several reasons
 - ▶ Simple
 - Explicit mechanism for calculating explicit probabilities for different hypothesis
 - Interpretable
- Initial probabilities
 - Domain knowledge
 - Learnt from training data
- Often gives good performance even when independence is violated in real world
- simple methods for parameter estimation, but is often not true

Assuming continuous attribute is Gaussian has nice analytical properties and yields a



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Extending NBC to Continuous Attributes

- Assuming attribute is Gaussian distribution
 Nice analytical properties
 - Simple method for parameter estimation
- But, is it sufficient?

E.g. Activity Start Time attribute for Eating Activity.

What do you expect the distribution to look like?





Extending NBC to Continuous Attributes

- Assuming attribute is Gaussian distribution
 Nice analytical properties
 - Simple method for parameter estimation
- But, is it sufficient?

E.g. Activity Start Time attribute for Eating Activity.

What do you expect the distribution to look like?





Gaussian Mixture Models



 Distribution P(Xi|Y=k) is modeled as a combination of M Gaussian probability distribution functions

$$P(X^{i}|Y = k) = \sum_{m=1}^{M} \pi_{m} N(\mu_{km}^{i}, \sigma_{km}^{2^{i}})$$

$$\pi_{m} \ge 0 \text{ and } \sum_{m=1}^{M} \pi_{m} = 1$$

• Generalize to cover all attributes by considering M multivariate Gaussians where $\mu_{km} \& \sigma_{km}^2$ are D-dimensional vectors

 Parameters and mixing coefficients estimated using an iterative Expectation-Maximization algorithm

Expectation Maximization Algorithm

Algorithm GMM_EM(x)

for m = 1 ... M

Initialize the multivariate Gaussian mean μ_{km} and variance σ_{km}^2 done

repeat for $i = 1 ... N_k$ $p_{im} = N(x_i \mid \mu_{km}, \sigma_{km}^2)$

 $z_i = \operatorname{argmax}_m p_{im}$

done

for m = 1 ... M A_m = number of data points assigned to mth component $\mu_{km} = \frac{1}{A_m} \sum \{x_i : z_i = m\}$ // Update the mean $\sigma_{km}^2 = \frac{1}{A_m} \sum \{ (x_i - \mu_{km})^2 : z_i = m \} // \text{ Update the variance}$ A_m = done until convergence

```
// x = x_1, x_2, ..., x_{Nk} is the subset L_k of the training data that is marked with activity k
// M is the number of Gaussian mixture components required to describe Lk
                                          // Expectation step
                                          // p_{im} represents probability that data point x_i has been
                                          // sampled from the m<sup>th</sup> mixture component N(\mu_{km}, \sigma_{kn}^2)
                                          // z_i is a latent variable representing the mixture
                                          // component to which data point x, is assigned
                                          // Maximization step
                                                 // Update the mixing coefficient
```


Expectation Maximization Algorithm

Initial Estimate

Algorithm GMM_EM(x)

// $x = x_1, x_2, ..., x_{Nk}$ is the subset L_k of the training data that is marked with activity k // M is the number of Gaussian mixture components required to describe Lk for m = 1 .. MInitialize the multivariate Gaussian mean μ_{km} and variance σ_{km}^2 done repeat for $i = 1 ... N_k$ // Expectation step $p_{im} = N(x_i \mid \mu_{km}, \sigma_{km}^2)$ // p_{im} represents probability that data point x_i has been // sampled from the mth mixture component $N(\mu_{km}, \sigma_{kn}^2)$ $// z_i$ is a latent variable representing the mixture $z_i = \operatorname{argmax}_m p_{im}$ // component to which data point x, is assigned done for m = 1 ... M// Maximization step A_m = number of data points assigned to mth component $\mu_{km} = \frac{1}{A_m} \sum \{x_i : z_i = m\}$ // Update the mean $\sigma_{km}^2 = \frac{1}{A_m} \sum \{(x_i - \mu_{km})^2 : z_i = m\} // \text{ Update the variance}$ A_m // Update the mixing coefficient = done until convergence

Cook & Krishnan. Activity Learning: Discovering, Recognizing, and Predicting Human Behavior from Sensor Data. 2015.



Expectation Maximization Algorithm



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Expectation Maximization Algorithm



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Some Other Classifiers: Decision Tree



Cook & Krishnan. Activity Learning: Discovering, Recognizing, and Predicting Human Behavior from Sensor Data. 2015.



Some Other Classifiers: Support Vector Machine



Cook & Krishnan. Activity Learning: Discovering, Recognizing, and Predicting Human Behavior from Sensor Data. 2015.



Some Other Classifiers: Support Vector Machine



Cook & Krishnan. Activity Learning: Discovering, Recognizing, and Predicting Human Behavior from Sensor Data. 2015.

Some Other Classifiers: Non-Linear Support Vector Machine



Non-separable Case

ar kernel	$x_i^T x_j$
nomial kernel	$(x_i^T x_j + c)^d$
al basis function kernel	$\exp\left(-\frac{ x_i - x_j _2^2}{2\sigma^2}\right)$
rbolic tangent kernel	$\tan h(\kappa x_i^T x_j + c)$ for $\kappa > 0$ and $c < 0$





SVM in practice...

Interesti	ng 🗸 GPUServer AWSUC	LA Frank McSherry	Home Network 🗸	Stanford CPNT	AWS Distill	>:	
Voice	Google Voic	Keras Docu	JupyterLab	Huma	SVM	-	
	Human Activity Recog Acknowledgements: This notebo • https://github.com/patoalejor	ook is based on material dra	awn from:	b/master/Recognition	.ipynb		
[18]:	<pre>import numpy as np import pandas as pd import time import pdb # usage: pdb.s import sys import matplotlib.pyplot %matplotlib inline</pre>						
	Download and read HAR Data Download data from Kaggle: https need to create a Kaggle account, Put them in ./data The data comes from recordings of waist-mounted smartphone with of WALKING_UPSTAIRS, WALKING Galaxy S II) on the waist. More de	and would get two files: tes of 30 study participants per embedded inertial sensors. _DOWNSTAIRS, SITTING, S	st.csv and train.csv forming activities of daily Each person performed s STANDING, LAYING) wear	living (ADL) while carr	ying a à,		
in [6]:]: df_test = pd.read_csv("data/test.csv") df_train = pd.read_csv("data/train.csv")						
	Explore and preprocess the dat	a					
n [24]:	<pre>print("Number of features print("Number of records print("Number of features print("Number of records trainData = df train.dro</pre>	<pre>in Train : ",train. in Test : ",test.s in Test : ",test.s</pre>	<pre>shape[0]) hape[1]) hape[0])</pre>	165			
	<pre>trainLabel = df_train.Act testData = df_test.drop(testLabel = df_test.Activ</pre>	<pre>ivity.values ['subject','Activity</pre>					
	<pre>print("Train Data shape print("Train Label shape print("Test Data shape print("Test Label shape</pre>	<pre>- : ",trainData.shape) : ",trainLabel.shape : ",testData.shape)</pre>)				
	<pre>print("Label examples: ") print(np.unique(trainLabe</pre>						
	Number of features in Tra Number of records in Tra Number of features in Tes Number of records in Tes Train Data shape : (735 Train Label shape : (735	in : 7352 t : 563 t : 2947 2, 561)					

Test Data shape : (2947, 561) Test Label shape : (2947,) Label examples: ['LAYING' 'SITTING' 'STANDING' 'WALKING' 'WALKING_DOWNSTAIRS' 'WALKING_UPSTAIRS']

Temporal Probabilistic Models

State Transition Model

- Specifies how the state of the environment evolves, i.e. $P(S_t | S_{0:t-1})$ Initial state is considered known $S_0 = s_0$
- Problem #1: the set $S_{0:t-1}$ is unbounded as t grows

 - Many flavors
 - simplest is first-order Markov process in which current state depends only on the previous state, i.e. $P(S_t | S_{0:t-1}) = P(S_t | S_{t-1})$



• Solve by Markov assumption: current state S_t depends only on a finite fixed number of previous states

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- - Solve by Stationarity Assumption, i.e. $P(S_t | S_{0:t-1})$

• Solve by Markov assumption: current state S_t depends only on a finite fixed number of previous states

$$) = \mathbf{P}(\mathbf{S}_{t'} | \mathbf{S}_{0:t'-1}) \forall t, t'$$





Sensor Model (aka Observation Model)

Russell, Stuart J., and Peter Norvig. "Artificial Intelligence: A Modern Approach"



Sensor Model (aka Observation Model)

- Specifies how the sensor observations depend on the current state and previous variables, i.e. $P(O_t | S_{0:t}, O_{1:t-1})$
- A good choice of state should suffice to generate the current sensor value, which leads to sensor Markov assumption, i.e. $P(O_t | S_{0:t}, O_{1:t-1}) = P(O_t | S_t)$



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- Combining with the simplified first order Markov transition model, we get:



• Complete joint distribution: $P(S_{0:t}, O_{1:t}) = P(S_0) \prod P(S_i | S_{i-1}) P(O_i | S_i)$

i=1

Dynamic Bayesian Network

Digression: Bayesian Network

- directed acyclic graph (DAG)

 - Edges represent conditional dependencies
- Efficient algorithms can perform inference and learning in Bayesian networks
 - Using a Bayesian network can save considerable amounts of memory over exhaustive probability tables,



• A probabilistic graphical model that represents a set of variables and their conditional dependencies via a

Nodes represent variables: observable quantities, latent variables, unknown parameters or hypotheses

- nodes that are not connected (no path connects one node to another) represent conditionally independent variables • Each node is associated with a probability function that takes, as input, a particular set of values for the node's parent variables, and gives the probability distribution of the variable represented by the node.

$P(G, S, R) = P(G \mid S, R)P(S \mid R)P(R)$

E.g. The model can answer questions about the presence of a cause given the presence of an effect (so-called inverse probability) like "What is the probability that it is raining, given the grass is wet?"

$$P(R = T \mid G = T) = \frac{P(G = T, R = T)}{P(G = T)} = \frac{\sum_{S \in \{T, F\}} P(G = T, S, R = T)}{\sum_{S, R \in \{T, F\}} P(G = T, S, R)}$$





Inference Tasks: Latent States from Sensor Observations



- Filtering: $\mathbf{P}(\mathbf{S}_t | \mathbf{O}_{1:t} = \mathbf{O}_{1:t})$
- Forecasting: $P(S_{t'} | O_{1:t} = O_{1:t})$ where t' > t
- Smoothing: $P(S_{t'} | O_{1:t} = O_{1:t})$ where t' < t
- Most likely explanation: $\operatorname{argmax}_{\mathbf{S}_{0:t}} \mathbf{P}(\mathbf{S}_{1:t} | \mathbf{O}_{1:t} = \mathbf{O}_{1:t})$

keep track of current state for rational decision-making

evaluate possible courses of action based on their expected outcomes

better estimate of the state than was available at the time, useful for learning

tasks such as speech recognition





Learning Task

- The transition and sensor models, if not yet known, can be learnt from observations
- Learning can be done as a byproduct of inference Inference provides an estimate of what transitions actually occurred and of what states generated the sensor readings
 - These estimates can be used to update the models
 - The updated model provides new estimates, and the process iterates to convergence Overall process is an instance of the Expectation Maximization (EM) algorithms



Filtering

- cost of update will increase with time
- In other words, we need a recursive estimation algorithm of the form $\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{o}_{1:t+1}) = f(\mathbf{o}_{t+1}, \mathbf{P}(\mathbf{S}_t | \mathbf{o}_{1:t}))$ for some function f

$$\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{o}_{1:t+1}) = \mathbf{P}(\mathbf{S}_{t+1} | \mathbf{o}_{1:t}, \mathbf{o}_{t+1})$$

= $\alpha \mathbf{P}(\mathbf{o}_{t+1} | \mathbf{S}_{t+1}, \mathbf{o}_{1:t}) \mathbf{P}(\mathbf{S}_{t+1} | \mathbf{o}_{1:t})$
= $\alpha \mathbf{P}(\mathbf{o}_{t+1} | \mathbf{S}_{t+1}) \mathbf{P}(\mathbf{S}_{t+1} | \mathbf{o}_{1:t})$
update prediction

 A useful filtering algorithm needs to maintain a current state estimate and update it In the rather than going back over the entire history of percepts for each update. as then the

• Computation can be viewed as being composed of two parts: project current state forward in time, and then update it in light of the new sensor observation O_{t+1}

> (dividing up the observations) (using Bayes' rule, α is normalizing constant) $|0_{1:t})$ (by the sensor Markov assumption)



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- forward in time, and then *update* it in light of the new sensor observation $\mathbf{0}_{t+1}$

$$\begin{aligned} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{o}_{1:t+1}) &= \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{o}_{1:t}, \mathbf{o}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}, \mathbf{o}_{1:t}) \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{o}_{1:t}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{o}_{1:t}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1}) \\ &= \alpha \mathbf{P}(\mathbf{e}_{t+1} \mid \mathbf{S}_{t+1}) \sum_{\mathbf{s}_{t}} \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{e}_{t+1})$$

Russell, Stuart J., and Peter Norvig. "Artificial Intelligence: A Modern Approach"

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(using Bayes' rule, α is normalizing constant) $(0_{1:t})$ (by the sensor Markov assumption)

 $\mathbf{s}_{t}, \mathbf{o}_{1:t}) \mathbf{P}(\mathbf{s}_{t} | \mathbf{o}_{1:t})$



(by Markov assumption)





Filtering

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$$\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{o}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{o}_{t+1} | \mathbf{S}_{t+1}) \sum_{\mathbf{s}_t}$$

above equation and $\mathbf{f}_{1:0} = \mathbf{P}(\mathbf{S}_0)$

• A useful filtering algorithm needs to maintain a current state estimate and update it ▶ rather than going back over the entire history of percepts for each update. as then the

 $\mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{s}_t) \mathbf{P}(\mathbf{s}_t \mid \mathbf{0}_{1:t})$

• One can view the filtered estimate $P(S_t | o_{1,t})$ as a message $f_{1,t}$ that is propagated along the sequence, modified by each transition and updated by each new sensor observation via $\mathbf{f}_{1:t+1} = FORWARD(\mathbf{f}_{1:t}, \mathbf{0}_{t+1})$ where FORWARD implements the

Example of Filtering



- Goal: compute $P(R_2 | u_{1.2})$
- Day 0: no observation only prior belief $P(R_0) = \langle 0.5, 0.5 \rangle$.
- Day 1: umbrella appears, so $U_1 = true$ $P(R_1) = \sum P(R_1 | r_0) P(r_0) = \langle 0.7, 0.3 \rangle \times 0.5 + \langle 0.3, 0.7 \rangle \times 0.5 = \langle 0.5, 0.5 \rangle$ $P(R_1 \mid u_1) = \alpha P(u_1 \mid R_1) P(R_1) = \alpha \langle 0.9, 0.2 \rangle \langle 0.5, 0.5 \rangle = \alpha \langle 0.45, 0.1 \rangle \approx \langle 0.818, 0.182 \rangle$
- Day 2: umbrella appears, so $U_2 = true$ $P(R_2 \mid u_1) = \sum P(R_2 \mid r_1) P(r_1 \mid u_1) = \langle 0.7, 0.3 \rangle \times 0.818 + \langle 0.3, 0.7 \rangle \times 0.182 \approx \langle 0.627, 0.373 \rangle$

$$P(R_2 | u_1, u_2) = \alpha P(u_2 | R_2) P(R_2 | u_1) = \alpha \langle 0.9, 0.2 \rangle \langle 0.627, 0.373 \rangle =$$



 $= \alpha \langle 0.565, 0.075 \rangle \approx \langle 0.883, 0.117 \rangle$

Intuitively, the probability of rain increases from day 1 to day 2 because rain persists.

Forecasting

- The filtering process already incorporates a one-step prediction
- It is easy to derive the following recursive computation for predicting the state at t + k + 1 from a prediction for t + k

$$P(\mathbf{S}_{t+k+1} | \mathbf{o}_{1:t}) = \sum_{\mathbf{s}_{t+k}} P(\mathbf{S}_{t+k+1} | \mathbf{s}_{t+k})$$
transition
model

- Note that no sensor model is involved in forecasting.
- to the stationary distribution of the Markov process defined by the transition model
 - **mixing time:** roughly, the time taken to reach the fixed point
 - the state space
 - obscured

• The task of forecasting can be seen simply as filtering without the addition of new sensor observation

 $\mathbf{s}_{t+k} \mathbf{P}(\mathbf{s}_{t+k} | \mathbf{o}_{1:t})$

recursion

• if we try to predict further and further into the future, the predicted distribution for the states will converge

- in practical terms, this dooms to failure any attempt to predict the actual state for a number of steps that is more than a small fraction of the mixing time, unless the stationary distribution itself is strongly peaked in a small area of

- the more uncertainty there is in the transition model, the shorter will be the mixing time and the more the future is





Smoothing



- Process of computing the distribution over past states given evidence up to the present i.e. $\mathbf{P}(\mathbf{S}_k | \mathbf{o}_{1:t})$ for $0 \le k < t$
- Another recursive message-passing approach by splitting the computation into two parts: the observations up to k and the observations from k + 1 to t

$$\begin{aligned} \mathbf{P}(\mathbf{S}_{k} | \mathbf{o}_{1:t}) &= \mathbf{P}(\mathbf{S}_{k} | \mathbf{o}_{1:k}, \mathbf{o}_{k+1,t}) \\ &= \alpha \mathbf{P}(\mathbf{S}_{k} | \mathbf{o}_{1:k}) \mathbf{P}(\mathbf{o}_{k+1,t} | \mathbf{S}_{k}, \mathbf{o}_{1:k}) \\ &= \alpha \mathbf{P}(\mathbf{S}_{k} | \mathbf{o}_{1:k}) \mathbf{P}(\mathbf{o}_{k+1,t} | \mathbf{S}_{k}) \\ &= \alpha \mathbf{f}_{1:k} \times \mathbf{b}_{k+1:1} \\ & \text{backward} \\ & \text{message} \end{aligned}$$

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(using Bayes' rule, given $\mathbf{o}_{1:k}$) (using conditional independence) $(\times \text{ represents pointwise multiplication of vectors})$



Computing Backward Message $b_{k+1:t}$

• By a recursive process that runs backward from t

$$\mathbf{P}(\mathbf{o}_{k+1:t} | \mathbf{S}_k) = \sum_{\mathbf{s}_{k+1}} \mathbf{P}(\mathbf{o}_{k+1:t} | \mathbf{S}_k, \mathbf{s}_{k+1}) \mathbf{P}(\mathbf{s}_{k+1})$$
$$= \sum_{\mathbf{s}_{k+1}} \mathbf{P}(\mathbf{o}_{k+1:t} | \mathbf{s}_{k+1}) \mathbf{P}(\mathbf{s}_{k+1} | \mathbf{S}_k)$$
$$= \sum_{\mathbf{s}_{k+1}} \mathbf{P}(\mathbf{o}_{k+1}, \mathbf{o}_{k+2:t} | \mathbf{s}_{k+1}) \mathbf{P}(\mathbf{s}_k)$$
$$= \sum_{\mathbf{s}_{k+1}} \mathbf{P}(\mathbf{o}_{k+1} | \mathbf{s}_{k+1}) \mathbf{P}(\mathbf{o}_{k+2:t} | \mathbf{s}_k)$$
recursion

- In message form, we have $\mathbf{b}_{k+1:t} = \mathsf{BACKWARD}(\mathbf{b}_{k+2:t}, \mathbf{o}_{k+1})$
- sequence and so the probability of observing it is 1.

(conditioning on S_{k+1}) $|\mathbf{S}_k)$ (by conditional independence) $_{k})$

 $_{k+1} | \mathbf{S}_k)$

(by conditional independence of \mathbf{o}_{k+1} & $\mathbf{o}_{k+2:t}$ given \mathbf{s}_{k+1} $_{k+1}$)**P**($\mathbf{s}_{k+1} \mid \mathbf{S}_k$) transition model

• Initialization with $\mathbf{b}_{t+1:t} = \mathbf{P}(\mathbf{0}_{t+1:t} | \mathbf{S}_t) = \mathbf{P}(|\mathbf{S}_t) = \mathbf{1}$ where 1 is a vector of 1s because $\mathbf{0}_{t+1:t}$ is an empty





Example of Smoothing



Goal: computing the smoothed estimate for the probability of rain at time k = 1 given the umbrella observations on days 1 and 2

$$P(R_{1} | u_{1}, u_{2}) = \alpha P(R_{1} | u_{1}) P(u_{2} | R_{1})$$

we know $P(R_{1} | u_{1}) = \langle 0.818, 0.182 \rangle$
while $P(u_{2} | R_{1})$ can be computed by applying the $P(u_{2} | R_{1}) = \sum_{r_{2}} P(u_{2} | r_{2}) P(|r_{2}| R_{1}) = (0.9 \times 1)$

so that $P(R_1 | u_1, u_2) = \alpha \langle 0.818, 0.182 \rangle \times \langle 0.69, 0.41 \rangle \approx \langle 0.883, 0.117 \rangle$

makes it more likely to have rained on day 1.

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he backward recursion

 $\times (0.7,0.3) + (0.2 \times 1 \times (0.3,0.7)) = (0.69,0.41)$

• Note that the smoothed estimate for rain on day 1 is higher than the filtered estimate (0.818) because the umbrella on day 2 makes it more likely to have rained on day 2; in turn, because rain tends to persist, that





Forward-Backward Algorithm for Smoothing

- Both the forward and backward recursions take a constant amount of time per step; hence, the time complexity of smoothing at time step k with respect to observation $\mathbf{0}_{1,t}$ is O(t)
- If we want to smooth the whole sequence, one obvious method is simply to run the whole smoothing process once for each time step to be smoothed. • This results in a time complexity of $O(t^2)$
- A better approach uses a simple application of dynamic programming to reduce the complexity to O(t)• The key to the linear-time algorithm is to record the results of forward filtering over the whole sequence • Then we run the backward recursion from t down to 1, computing the smoothed estimate at each step k from the computed backward message $\mathbf{b}_{k+1:t}$ and stored forward message $\mathbf{f}_{1:k}$

function FORWARD-BACKWARD(ev, prior) returns a vector of probability distributions inputs: ev, a vector of evidence values for steps 1,...,t *prior*, the prior distribution on the initial state, $P(S_0)$ **local variables:** fv, a vector of forward messages for steps $0, \ldots, t$ sv, a vector of smoothed estimates for steps $1, \ldots, t$

```
\mathbf{fv}[0] \leftarrow prior
for i = 1 to t do
       \mathbf{fv}[i] \leftarrow \text{FORWARD}(\mathbf{fv}[i-1], \mathbf{ev}[i])
for i = t down to 1 do
       \mathbf{sv}[i] \leftarrow \text{NORMALIZE}(\mathbf{fv}[i] \times \mathbf{b})
       \mathbf{b} \leftarrow \text{BACKWARD}(\mathbf{b}, \mathbf{ev}[i])
return sv
```

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b, a representation of the backward message, initially all 1s



Drawbacks of Forward-Backward Algorithm in Practice

- Space complexity can be too high when the state space is large and the sequences are long. • It uses $O(|\mathbf{f}|t)$ space where $|\mathbf{f}|$ is the size of the representation of the forward message • Can be reduced to $O(|\mathbf{f}| \log t)$ with a concomitant increase in the time complexity by a factor of log t
- It needs to be modified to work in an online setting where smoothed estimates must be computed for earlier time slices as new observations are continuously added to the end of the sequence • The most common requirement is for fixed-lag smoothing, which requires computing the smoothed estimate $\mathbf{P}(\mathbf{S}_{t-d} | \mathbf{o}_{1:t})$ for fixed d,

i.e. smoothing is done for time step that is d steps behind the current time

- Inefficient to run Forward-Backward over the d-step window for each sensor observation
- \bullet Instead, fixed-lag smoothing can, in some cases, be done in constant time per update independent of d

Finding the Most Likely Sequence

- Brute force approach enumerate all state sequences and computer their likelihood: expensive!
- Incorrect approach
 - using at each step the state that is most likely according to the posterior
 - most likely sequence \neq sequence of most likely states
 - likely sequence we must consider *joint* probabilities over all the time steps.
- Efficient linear time (and space) algorithm: the Viterbi algorithm

• Problem: given sensor observation sequence $\mathbf{0}_{1,t}$ what state sequence is most likely to explain this?

• use smoothing to find the posterior distribution for the state at each time step; then construct the sequence,

- posterior distributions computed by smoothing are distributions over *single* time steps, whereas to find the most



Viterbi Algorithm

- the given observations at each state
- \mathbf{S}_{t+1} and most likely paths to each state \mathbf{S}_t
 - Most likely path to each $s_{t+1} = most$ likely path to some s_t plus one more step

 $\max_{\mathbf{s}_{1:t}} \mathbf{P}(\mathbf{s}_{1:t}, \mathbf{S}_{t+1} | \mathbf{o}_{1:t+1}) = \mathbf{P}(\mathbf{o}_{t+1} | \mathbf{S}_{t+1}) \max_{\mathbf{s}_{t}} (\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{S}_{t+1} | \mathbf{S}_{t+1}))$

or, $\mathbf{m}_{1:t+1} = \mathbf{P}(\mathbf{o}_{t+1} | \mathbf{S}_{t+1}) \max_{\mathbf{S}_t} (\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{S}_t) \mathbf{m}_{1:t})$

- Similar to filtering algorithm: it starts at time 0 and and then runs forward along the sequence, computing the vector message **m**
- - One can thus easily select the final state of the most likely sequence overall
 - record, for each state, the best state that leads to it

• View each state sequence as a path through a graph whose nodes are the possible states at each time step • The likelihood of any path is the product of the transition probabilities along the path and the probabilities of

• Because of the Markov property, there is a recursive relationship between most likely paths to each state

$$\mathbf{S}_{t+1} | \mathbf{s}_t \max_{\mathbf{s}_{1:t-1}} \mathbf{P}(\mathbf{s}_{1:t-1}, \mathbf{s}_t | \mathbf{o}_{1:t}))$$
where $\mathbf{m}_{1:t} = \max_{\mathbf{s}_{1:t-1}} \mathbf{P}(\mathbf{s}_{1:t-1}, \mathbf{S}_t | \mathbf{o}_{1:t}))$, and $\mathbf{m}_{1:0} = \mathbf{P}(\mathbf{S}_0)$
and then runs forward along the sequence, computing the

• Eventually $\mathbf{m}_{1,t}$ will contain the probability for the most likely sequence reaching each of the final states

- In order to identify the actual sequence, as opposed to just computing its probability, the algorithm will also need to





Example of Most Likely Sequence





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Hidden Markov Models

- A temporal probabilistic model in which the process state is described by a single, discrete RV The possible values of the variable are the possible states of the world
- The restricted structure allows for a simple and elegant matrix implementation of all the basic algorithms • Let the state variable be an integer $S_t \in \{1, 2, \dots, S\}$ where S is the number of possible states
- - Fransition model $\mathbf{P}(S_t | S_{t-1})$ becomes an $S \times S$ matrix \mathbf{T} where $\mathbf{T}_{ii} = P(S_t = j | S_{t-1} = i)$
 - Rain/Umbrella example

$$\mathbf{T} = \mathbf{P}(S_t = S_{t-1}) = \begin{pmatrix} 0.7 & 0.3 \\ 0.3 & 0.7 \end{pmatrix}$$

- What happens if one has a model with two or more state variables?
 - One can still fit it into the HMM framework by combining the variables into a single "megavariable" whose values are all possible tuples of values of the individual state variables.



Hidden Markov Models (contd.)

- Although the state must be a single, discrete variable, there is no similar restriction on observations. This is because observations variations are known, so that there is no need to keep track their distributions ▶ If a variable is not observed, it can simply be dropped from the model for that time step. There can be many observation variables, both discrete and continuous
- However, one can also put sensor model in matrix form
 - occur, i.e. $P(\mathbf{0}_t | S_t = i) \quad \forall i \in \{1, 2, \dots S\}$
 - Specified as $S \times S$ diagonal matrix \mathbf{O}_t where the *i*-th diagonal entry is $\mathbf{P}(\mathbf{O}_t | S_t = i)$
- Umbrella world:
 - If $Umbrella_t = true$ then

$$\mathbf{O}_t = \begin{pmatrix} 0.9 & 0\\ 0 & 0.2 \end{pmatrix}$$

• and if $Umbrella_t = false$ then

$$\mathbf{O}_t = \begin{pmatrix} 0.1 & 0\\ 0 & 0.8 \end{pmatrix}$$

Russell, Stuart J., and Peter Norvig. "Artificial Intelligence: A Modern Approach"

• Because observation $\mathbf{0}_{t}$ at every time step t is known, only need to specify for each state how likely is $\mathbf{0}_{t}$ to

Forward-Backward Algorithm in Matrix-Vector Form

- Column vectors to represent the forward & backward messages
 - forward equation becomes:

$$\mathbf{f}_{1:t+1} = \alpha \mathbf{O}_{t+1} \mathbf{T}^{\mathsf{T}} \mathbf{f}_{1:t}$$

backward equation becomes:

$$\mathbf{b}_{k+1:t} = \mathbf{TO}_{k+1}\mathbf{b}_{k+2:t}$$

- Complexity
 - Time: $O(S^2t)$
 - each step requires multiplying an S element vector with an $S \times S$ matrix
 - Space: O(St)
 - the forward pass stores stores t vectors of size S
- Matrix-Vector formulation allows for improvements too, e.g.,
 - constant space algorithm for smoothing, independent of t
 - online smoothing with fixed lag

function FORWARD-BACKWARD(ev, prior) returns a vector of probability distributions inputs: ev, a vector of evidence values for steps 1,...,t *prior*, the prior distribution on the initial state, $P(S_0)$ local variables: fv, a vector of forward messages for steps $0, \ldots, t$ **b**, a representation of the backward message, initially all 1s sv, a vector of smoothed estimates for steps $1, \ldots, t$

```
\mathbf{fv}[0] \leftarrow prior
for i = 1 to t do
       \mathbf{fv}[i] \leftarrow FORWARD(\mathbf{fv}[i-1], \mathbf{ev}[i])
for i = t down to 1 do
       \mathbf{sv}[i] \leftarrow \text{NORMALIZE}(\mathbf{fv}[i] \times \mathbf{b})
       \mathbf{b} \leftarrow \text{BACKWARD}(\mathbf{b}, \mathbf{ev}[i])
return sv
```





Constant Space Algorithm for Smoothing (independent of t)

- Idea: Smoothing for any particular time index k requires the simultaneous presence of both the forward and backward messages $\mathbf{f}_{1:k}$ and $\mathbf{b}_{k+1:t}$ $P(S_k | \mathbf{o}_{1:t}) = \alpha \mathbf{f}_{1:k} \times \mathbf{b}_{k+1:1}$
- Forward-Backward algorithm achieves this by storing the fs computed on the forward pass so that they are available during the backward pass.
- Instead, one can do a single pass that propagates both **f** and **b** in the same direction
 First do standard forward pass to compute **f**_{t:t}
 forgetting all the intermediate results
 Then run backward pass for both **f** and **b** together using a backward propagation of **f** as
 - Then run backward pass for both **f** and **b** together using a backward propagation of **f** as $\mathbf{f}_{1:t} = \alpha'(\mathbf{T}^{\mathsf{T}})^{-1}\mathbf{O}_{t+1}^{-1}\mathbf{f}_{1:t+1}$



Online Smoothing with Fixed Lag d

- How to do this incrementally? First, we can compute $\mathbf{f}_{1:t-d+1}$ from $\mathbf{f}_{1:t-d}$ using the standard filtering process $P(S_{t+1} | \mathbf{o}_{1:t+1}) = \alpha P(\mathbf{o}_{t+1} | S_{t+1}) \sum P(S_{t+1} | s_t) P(s_t | \mathbf{o}_{1:t})$
 - Trick: repeatedly use $\mathbf{b}_{k+1:t} = \mathbf{TO}_{k+1}\mathbf{b}_{k+2:t}$ to get i = t - d + 1

which gives the incremental update $\mathbf{B}_{t-d+2:t+1} = \mathbf{O}_{t-d+1}^{-1} \mathbf{T}^{-1} \mathbf{B}_{t-d+1:t} \mathbf{T} \mathbf{O}_{t+1}$

• Goal: smooth at time t - d when the current time is t, i.e. compute $\alpha \mathbf{f}_{1:t-d} \times \mathbf{b}_{t-d+1:1}$ • Then, when a new observation arrives, we need to compute $\alpha \mathbf{f}_{1:t-d+1} \times \mathbf{b}_{t-d+2:1}$

• But, there is no analogous simple relationship between new $b_{t-d+2:1}$ and old $b_{t-d+1:1}$

 $\mathbf{b}_{t-d+1:1} = (\prod_{i=1}^{r} \mathbf{TO}_{i})\mathbf{b}_{t+1:t} = \mathbf{B}_{t-d+1:t}\mathbf{1}$ and $\mathbf{b}_{t-d+2:1} = (\prod_{i=1}^{r} \mathbf{TO}_{i})\mathbf{b}_{t+2:t+1} = \mathbf{B}_{t-d+2:t+1}\mathbf{1}$ i = t - d + 2



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Online Algorithm for Smoothing with Fixed Lag d

inputs: e_t , the current evidence for time step t d, the length of the lag for smoothing **persistent**: t, the current time, initially 1 add e_t to the end of $e_{t-d:t}$ $\mathbf{O}_t \leftarrow \text{diagonal matrix containing } \mathbf{P}(e_t | X_t)$ if t > d then $\mathbf{f} \leftarrow \text{FORWARD}(\mathbf{f}, e_{t-d})$ remove e_{t-d-1} from the beginning of $e_{t-d:t}$ $\mathbf{O}_{t-d} \leftarrow \text{diagonal matrix containing } \mathbf{P}(e_{t-d} | X_{t-d})$ $\mathbf{B} \leftarrow \mathbf{O}_{t-d}^{-1} \mathbf{T}^{-1} \mathbf{B} \mathbf{T} \mathbf{O}_t$ else $\mathbf{B} \leftarrow \mathbf{BTO}_t$ $t \leftarrow t+1$ if t > d + 1 then return NORMALIZE($\mathbf{f} \times \mathbf{B1}$) else return null

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```
function FIXED-LAG-SMOOTHING(e_t, hmm, d) returns a distribution over \mathbf{X}_{t-d}
           hmm, a hidden Markov model with S \times S transition matrix T
               f, the forward message \mathbf{P}(X_t | e_{1:t}), initially hmm.PRIOR
               B, the d-step backward transformation matrix, initially the identity matrix
               e_{t-d;t}, double-ended list of evidence from t - d to t, initially empty
  local variables: O_{t-d}, O_t, diagonal matrices containing the sensor model information
```



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HMM Application: Vacuum Robot Localization



- Discrete grid where $s_t \in 1, 2, 3, \dots, 42$ represents the current location of the robot
- Robot can move randomly to any of the adjacent empty squares from its current location i.e., $P(S_{t+1} = j | S_t = i) = T_{ii} = 1/(\# \text{ of empty adjacent locations})$ if $j \in \text{NEIGHBORS}(i)$ else 0
- Noisy binary sensors with error rate ϵ report whether the four adjacent NESW locations are empty or not, i.e. O_t is a 4-bit sequence taking values such as $o_t = 1011$ so that $P(O_t = o_t | S_t = i) = (\mathbf{0}_t)_{ii} = (1 - \epsilon)^{4 - d_{it}} \epsilon^{d_{it}}$ d_{it} = is the discrepancy between the true values for square *i* and the actual reading o_t where The robot can use filtering to estimate current location, and smoothing to get past location



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 Consider vacuum robot that has the policy of going straight for as long as it can $42 \times 8 = 168$ values, so that **T** will be of size $168^2 = 28,224$ - still manageable though large

Then the state will need to including both *location* and *heading*, and so could take



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- What if we add the possibility of dirt in each of the 42 squares?

• Now the number of states is multiplied by 2^{42} and ${f T}$ has more than 10^{29} entries ${f C}$



- Consider vacuum robot that has the policy of going straight for as long as it can Then the state will need to including both *location* and *heading*, and so could take $42 \times 8 = 168$ values, so that **T** will be of size $168^2 = 28,224$ - still manageable though large
- What if we add the possibility of dirt in each of the 42 squares? • Now the number of states is multiplied by 2^{42} and ${f T}$ has more than 10^{29} entries ${f C}$ • In general if state is composed of n discrete variables that can take at most d values
- each, then the state transition matrix will have size $O(d^{2n})$ and the per-update computation time will be also $O(d^{2n})$
 - ▶ While HMMs used a lot (e.g. speech recognition), they are fundamentally limited in their ability to represent complex processes Problem: represent states as integers and don't exploit any internal structure



Kalman Filter



- sugar level, etc.) from irregular and noisy observations
- If variables are discrete, one can use HMMs
- Kalman Filter targets the case when the variables of interest are continuous three for position (X_t, Y_t, Z_t) and three for velocity $(\dot{X}_t, \dot{Y}_t, \dot{Z}_t)$
- next state S_{t+1} is a linear function of current state S_t plus some Gaussian noise
 - sensor observation O_t is a linear function of current state S_t plus some Gaussian noise



In many applications we need to reconstruct trajectory of some variable of interest (position, speed, blood

• E.g. flight of incoming enemy aircraft or a pedestrian on the road may be specified by six continuous variables:

• Uses linear-Gaussian distributions for conditional densities to represent the transition and sensor models





Consider Tracking a Vehicle from Sporadic Observations

•
$$\mathbf{S}_t = (X_t, Y_t, Z_t, \dot{X}_t, \dot{Y}_t, \dot{Z}_t)$$

- Consider just X_t coordinate and let Δ be the time interval between observations
- Then, assuming constant velocity during the interval and no noise $X_{t+\Delta} = X_t + \dot{X}_t \Delta$
- Adding Gaussian noise (to account for wind variation, etc.), we obtain a linear–Gaussian transition model: $P(X_{t+\Delta} = x_{t+\Delta} | X_t = x_t, \dot{X}_t = \dot{x}_t) = N(x_{t+\Delta}; x_t + \dot{x}_t\Delta, \sigma^2)$
- Note: this is a specific case of linear-Gaussian, and when generalizing to d dimensions then we need to use multivariate Gaussian with a d-element mean and a $d\times d$ covariance matrix Σ





Updating Gaussian Distributions

 Special property of linear-Gaussian: closed under Bayesian updating ▶ i.e. given an observation, the posterior distribution is still in the linear–Gaussian family

• Recall filtering:
$$\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{o}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{o}_{t+1} | \mathbf{S}_{t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{S}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{S}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) \sum_{\substack{s \in \mathbf{S}_t \\ \text{model}}} \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{e}_{1:t+1} | \mathbf{e}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{P}(\mathbf{e}_{1:t+1})$$

the the one-step predicted distribution given by the following is also Gaussian.

$$\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{o}_{1:t}) = \int_{\mathbf{S}_t} \mathbf{P}(\mathbf{S}_{t+1} | \mathbf{s}_t) \mathbf{P}(\mathbf{s}_t | \mathbf{o}_{1:t}) d\mathbf{s}_t$$

after conditioning on the new observation $\mathbf{0}_{t+1}$, the updated distribution below is also Gaussian.

$$\mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{o}_{1:t+1}) = \alpha \mathbf{P}(\mathbf{o}_{t+1} \mid \mathbf{S}_{t+1}) \mathbf{P}(\mathbf{S}_{t+1} \mid \mathbf{o}_{1:t+1})$$



• If current distribution $\mathbf{P}(\mathbf{S}_t | \mathbf{o}_{1:t})$ is Gaussian and the transition model $\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{s}_t)$ is linear-Gaussian then

• If the prediction $P(S_{t+1} | o_{1:t})$ is Gaussian and the sensor model $P(o_{t+1} | S_{t+1})$ is linear–Gaussian, then,

1:t/

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Kalman Filter in Summary

- mean μ_{t+1} and covariance Σ_{t+1} .
- produces a Gaussian state distribution for all time
 - grows without bound over time
- Simple univariate case: state variable X_t with nos
- General multivariate case

 $P(\mathbf{x}_{t+1}|\mathbf{x}_t) \hspace{0.1 in} = \hspace{0.1 in} N(\mathbf{x}_{t+1};$ $P(\mathbf{z}_t | \mathbf{x}_t) = N(\mathbf{z}_t; \mathbf{H})$ $\mu_{t+1} \hspace{.1in} = \hspace{.1in} \mathbf{F} \mu_t + \mathbf{K}_{t+1}(\mathbf{z}_{t+1})$ $\Sigma_{t+1} = (\mathbf{I} - \mathbf{K}_{t+1}\mathbf{H})(\mathbf{H})$

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• FORWARD operator for Kalman filtering takes a Gaussian forward message $\mathbf{f}_{1:t}$ specified by a mean μ_t and covariance Σ_t , and produces a new multivariate Gaussian forward message $f_{1,t+1}$ specified by a

This translates into computing a new mean and covariance from the previous mean and covariance

• So if we start with a Gaussian prior $\mathbf{f}_{1\cdot 0} = \mathbf{P}(\mathbf{S}_0) = N(\mu_0, \Sigma_0)$, filtering with a linear–Gaussian model

Not only nice and elegant, but also important: except for a few special cases such as this, filtering with continuous or hybrid (discrete and continuous) networks generates state distributions whose representation

e variable
$$X_t$$
 with nosy observation Z_t
 $\mu_{t+1} = \frac{(\sigma_t^2 + \sigma_x^2)z_{t+1} + \sigma_z^2\mu_t}{\sigma_t^2 + \sigma_x^2 + \sigma_z^2}$
 $\sigma_{t+1}^2 = \frac{(\sigma_t^2 + \sigma_x^2)z_{t+1} + \sigma_z^2\mu_t}{\sigma_t^2 + \sigma_z^2}$
 $P(\mathbf{x}_{t+1}|\mathbf{x}_t) = N(\mathbf{x}_{t+1}; \mathbf{F}\mathbf{x}_t, \Sigma_x),$
 $\mu_{t+1} = \mathbf{F}\mu_t + \mathbf{K}_{t+1}(\mathbf{z}_{t+1} - \mathbf{H}\mathbf{F}\mu_t),$
 $\Sigma_{t+1} = (\mathbf{I} - \mathbf{K}_{t+1}\mathbf{H})(\mathbf{F}\Sigma_t\mathbf{F}^\top + \Sigma_x),$
where $\mathbf{K}_{t+1} = (\mathbf{F}\Sigma_t\mathbf{F}^\top + \Sigma_x)\mathbf{H}^\top (\mathbf{H}(\mathbf{F}\Sigma_t\mathbf{F}^\top + \Sigma_x)\mathbf{H}^\top + \Sigma_z)^{-1}$ Kalman gain matrix





 $egin{aligned} P(\mathbf{x}_{t+1} | \mathbf{x}_t) &= N(\mathbf{x}_{t+1}; \mathbf{F}\mathbf{x}_t, \mathbf{X}_t, \mathbf{Y}_t) \ P(\mathbf{z}_t | \mathbf{x}_t) &= N(\mathbf{z}_t; \mathbf{H}\mathbf{x}_t, \mathbf{\Sigma}_t, \mathbf{X}_t) \ \mu_{t+1} &= \mathbf{F}\mu_t + \mathbf{K}_{t+1}(\mathbf{z}_{t+1} - \mathbf{H}_t) \ \Sigma_{t+1} &= (\mathbf{I} - \mathbf{K}_{t+1}\mathbf{H})(\mathbf{F}\mathbf{\Sigma}_t\mathbf{F}^{ op}) \ \mathbf{Where} \ \mathbf{K}_{t+1} &= (\mathbf{F}\mathbf{\Sigma}_t\mathbf{F}^{ op} + \mathbf{\Sigma}_x) \mathbf{I} \end{aligned}$

$$egin{aligned} & \Sigma_t, \Sigma_x, \ & - \mathbf{H} \mathbf{F} \mu_t, \ & \mathbf{F}^ op + \Sigma_x, \ & \mathbf{F}^ op + \Sigma_x, \ & \mathbf{F}^ op + \Sigma_x, \ & \mathbf{H}^ op (\mathbf{H} (\mathbf{F} \Sigma_t \mathbf{F}^ op + \Sigma_x) \mathbf{H}^ op + \Sigma_z)^{-1} & ext{Kalman gain matrix} \end{aligned}$$







$$egin{aligned} & \Sigma_t, \Sigma_x, \ & - \mathbf{H} \mathbf{F} \mu_t, \ & \mathbf{F}^ op + \Sigma_x, \ & \mathbf{F}^ op + \Sigma_x, \ & \mathbf{F}^ op + \Sigma_x, \ & \mathbf{H}^ op (\mathbf{H} (\mathbf{F} \Sigma_t \mathbf{F}^ op + \Sigma_x) \mathbf{H}^ op + \Sigma_z)^{-1} & ext{Kalman gain matrix} \end{aligned}$$























Real-time Computation of Kalman Filter

 $P(\mathbf{x}_{t+1}|\mathbf{x}_t) = N(\mathbf{x}_{t+1};\mathbf{F}\mathbf{x}_t,\Sigma_x)$ $P(\mathbf{z}_t | \mathbf{x}_t) = N(\mathbf{z}_t; \mathbf{H}\mathbf{x}_t, \Sigma_z),$ $\mu_{t+1} = \mathbf{F} \mu_t + \mathbf{K}_{t+1} (\mathbf{z}_{t+1} - \mathbf{H} \mathbf{F} \mu_t)$ $\Sigma_{t+1} = (\mathbf{I} - \mathbf{K}_{t+1}\mathbf{H})(\mathbf{F}\Sigma_t\mathbf{F}^\top + \Sigma_x)$

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where $\mathbf{K}_{t+1} = (\mathbf{F}\Sigma_t \mathbf{F}^\top + \Sigma_x) \mathbf{H}^\top (\mathbf{H}(\mathbf{F}\Sigma_t \mathbf{F}^\top + \Sigma_x) \mathbf{H}^\top + \Sigma_z)^{-1}$ Kalman gain matrix





Real-time Computation of Kalman Filter

$$egin{aligned} P(\mathbf{x}_{t+1} | \mathbf{x}_t) &= N(\mathbf{x}_{t+1}; \mathbf{F}\mathbf{x}_t) \ P(\mathbf{z}_t | \mathbf{x}_t) &= N(\mathbf{z}_t; \mathbf{H}\mathbf{x}_t, \ \mu_{t+1} &= \mathbf{F}\mu_t + \mathbf{K}_{t+1}(\mathbf{z}_{t+1} + \mathbf{X}_t) \ \Sigma_{t+1} &= (\mathbf{I} - \mathbf{K}_{t+1}\mathbf{H})(\mathbf{F}\Sigma_t) \ \mathbf{W} \text{here } \mathbf{K}_{t+1} = (\mathbf{F}\Sigma_t\mathbf{F}^ op + \Sigma_t) \end{aligned}$$







Real-time Computation of Kalman Filter

$$egin{aligned} P(\mathbf{x}_{t+1} | \mathbf{x}_t) &= N(\mathbf{x}_{t+1}; \mathbf{F}\mathbf{x}_t) \ P(\mathbf{z}_t | \mathbf{x}_t) &= N(\mathbf{z}_t; \mathbf{H}\mathbf{x}_t, \ \mu_{t+1} &= \mathbf{F}\mu_t + \mathbf{K}_{t+1}(\mathbf{z}_{t+1} + \mathbf{X}_t) \ \Sigma_{t+1} &= (\mathbf{I} - \mathbf{K}_{t+1}\mathbf{H})(\mathbf{F}\Sigma_t) \ \mathbf{W} \text{here } \mathbf{K}_{t+1} = (\mathbf{F}\Sigma_t\mathbf{F}^ op + \Sigma_t) \end{aligned}$$

The sequence of values for Σ_t and \mathbf{K}_t can be computed offline, and so calculations during online tracking are modest.







Kalman Filtering and Smoothing for Object Moving in X - Y Plane



- • • State variables are $\mathbf{X} = (X, Y, X, Y)$ so that $\mathbf{F}, \Sigma_x, \mathbf{H}$, and Σ_z are 4×4 matrices

• Variance in the right plot (Kalman Smoothing) is much reduced except at the start and the end. Why?



Applicability of Kalman Filtering

- Classical applications
 - vehicles and people
- More esoteric applications
 - measurements
- Applicable to any system with continuous state variables and noisy measurements • national economies, human body, chemical plants, nuclear reactors, plant ecosystems, etc.
- But results may not be valid or useful strong assumption of linear-Gaussian transition and sensor models
- Extended Kalman Filter (EKF): overcome nonlinearities in the system being modeled
 - models the system as locally linear in \mathbf{x}_t in the neighborhood of $\mathbf{x}_t = \mu_t$
 - works well for smooth, well-behaved systems
 - true posterior
- What about systems that are "unsmooth" and "poorly behaved"?

radar tracking of aircrafts and missiles, acoustic tracking of submarines and ground vehicles, visual tracking of

reconstruct particle trajectories from bubble chamber photographs, ocean currents from satellite surface

- allows the tracker to maintain and update a Gaussian state distribution that is a reasonable approximation to the





Switching Kalman Filter

- bird, and the mean of this Gaussian will be centered on the trunk
- - relative to the trunk



- - E.g., one for straight flight, one for sharp left turns, and one for sharp right turns

 - A special case of the general dynamic Bayesian network model

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• Track a bird as it flies through the jungle and appears to be heading at high speed straight for a tree trunk • The Kalman filter, whether regular or extended, can make only a Gaussian prediction of the location of the

• A reasonable model of the bird, on the other hand, would predict evasive action to one side or the other Such a model is highly nonlinear, because the bird's decision varies sharply depending on its precise location

Switching Kalman Filter: multiple Kalman filters run in parallel, each using a different model of the system

• A weighted sum of predictions is used, where the weight depends on how well each filter fits the current data



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Dynamic Bayesian Networks (DBNs)

• Each slice can have any number of state variables S_t and observation variables O_t



• For simplicity, we assume that slice to slice (*time-homogeneous*) only in its own slice or the immediately preceding slice.

Special case of Bayesian Networks: infinitely many variables grouped in time slices

It the variables, their links, and their conditional distributions are exactly replicated from

▶ the DBN represents a first-order Markov process, i.e., each variable can have parents







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• Every hidden Markov model can be represented as a DBN with a single state variable and a single evidence variable



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- Every hidden Markov model can be represented as a DBN with a single state variable and a single evidence variable
- Every discrete-variable DBN can be represented as an HMM
 - tuples of values of the individual state variables

• We can combine all the state variables in the DBN into a single state variable whose values are all possible



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- Every hidden Markov model can be represented as a DBN with a single state variable and a single evidence variable
- Every discrete-variable DBN can be represented as an HMM
 - tuples of values of the individual state variables
- If every HMM is a DBN and every DBN can be translated into an HMM, what's the difference?

• We can combine all the state variables in the DBN into a single state variable whose values are all possible



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- Every hidden Markov model can be represented as a DBN with a single state variable and a single evidence variable
- Every discrete-variable DBN can be represented as an HMM
 - tuples of values of the individual state variables
- If every HMM is a DBN and every DBN can be translated into an HMM, what's the difference?
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- advantage of sparseness in the temporal probability model.
- Consider: a temporal process with *n* discrete variables, each with up to *d* values • HMM model will need a transition matrix of size $O(d^{2n})$
 - DBN model has size $O(nd^k)$, assuming each node can have at most k parents

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• The difference: by decomposing the state of a complex system into its constituent variables, we can take

• E.g. vacuum robot with 42 possibly dirty locations: size reduces from 5×10^{29} to a few thousands



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conditional distributions

• Every Kalman filter model can be represented in a DBN with continuous variables and linear–Gaussian



- conditional distributions
- However, not every DBN can be represented by a Kalman filter model
 - ▶ In a Kalman filter, the current state distribution is always a single multivariate Gaussian distribution - e.g. a single "bump" in a particular location in a position tracking system
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 - locked in the car, ...
 - being in mid-air above the front garden!

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• They might be in my pocket, on the bedside table, on the kitchen counter, dangling from the front door, or

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- combinations of discrete and continuous variables in order to get reasonable models

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• A single Gaussian bump that included all these places would have to allocate significant probability to the keys

• Real world has purposive agents, obstacles, and pockets which introduce "nonlinearities" and require





Constructing DBNs

- To construct a DBN, one must specify three kinds of information: the prior distribution over the state variables, $P(S_0)$; the transition model, $P(S_{t+1} | S_t)$; and the sensor model, $P(O_t | S_t)$
- To specify the transition and sensor models, one must also specify the topology of the connections between successive slices and between the state and evidence variables.
- Because the transition and sensor models are assumed to be time-homogeneous—the same for all *t*—one only needs to specify $\mathbf{P}(\mathbf{S}_{t+1} | \mathbf{S}_t)$ and $\mathbf{P}(\mathbf{O}_t | \mathbf{S}_t)$ for the first slice
 - the complete DBN with an unbounded number of time slices can be constructed as needed by copying the first slice



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Example: Battery Powered Robot Moving in X - Y Plane

- State Variables
 - Position: $\mathbf{X}_t = (X_t, Y_t)$
 - Velocity: $\dot{\mathbf{X}}_t = (\dot{X}_t, \dot{Y}_t)$
 - ► Battery Status: *Battery*
- Topology:
 - The position at the next time step depends on the current position and velocity
 - The velocity at the next step depends on the current velocity and the state of the battery
 - The battery state at the next step depends on the current velocity and the current state of the battery
- Sensor observations
 - Some method of measuring position—perhaps a fixed camera or onboard GPS (Global Positioning System)—yielding \mathbf{Z}_{t}
 - ▶ Battery charge meter yielding *BMeter*,







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A Deeper Look at the Sensor Model $BMeter_t$

- Assume discrete values in the range 0 to 5 for both $Battery_t$ and $BMeter_t$
 - have probabilities of 1.0 along the diagonal and 0 elsewhere
- Reality: noise always creeps into measurements
 - ▶ For continuous measurements, a Gaussian distribution with a small variance might be used
 - ▶ For our discrete variables, we can approximate a Gaussian
 - using a distribution in which the probability of error drops off in the appropriate way, so that the probability of a large error is very small
- However, in real world sensors also *fail*
 - When a sensor fails, it simply sends nonsense
 - **Transient failure**: the sensor occasionally decides to send nonsense
 - model by a higher probability for large errors than gaussian noise
 - Persistent failure: the sensor fails and stays failed
 - model by an additional state variable *BMBroken*
 - Other issues: sensor drift, sudden decalibration, and the effects of exogenous conditions (such as weather) on sensor readings

• If the meter was always accurate then the conditional probability table (CPT) for $\mathbf{P}(BMeter_t | Battery_t)$ will

B_0	$P(B_1)$
t	1.000
f	0.001




Russell, Stuart J., and Peter Norvig. "Artificial Intelligence: A Modern Approach"



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by replicating slices until the network is large enough to accommodate the observations



Given a sequence of observations, one can construct the full Bayesian network representation of a DBN



84

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- But very inefficient 😰
 - Space requires to filter or smooth on $\mathbf{0}_{1:t}$ would require O(t) space and thus grown without bound
 - Inference time per update will also increase as O(t)

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- But very inefficient 😰
 - Space requires to filter or smooth on $\mathbf{0}_{1:t}$ would require O(t) space and thus grown without bound
 - Inference time per update will also increase as O(t)
- What about recursive algorithms with "constant" time and space complexity per filtering update?
- Unfortunately exponential: pace complexity $O(d^{n+k})$ and time complexity $O(nd^{n+k})$ with n state variables of domain size d and maximum number of parents k

Russell, Stuart J., and Peter Norvig. "Artificial Intelligence: A Modern Approach"

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Approximate inference in DBNs

- Idea: randomized sampling algorithms, aka Monte Carlo algorithms
 - answers found in those random events
 - Provide approximate answers whose accuracy depends on the number of samples generated
- Two families of algorithms used for general Bayes Networks
 - Direct sampling
 - generate each sample from scratch
 - e.g. Likelihood Weighting
 - Markov chain sampling
 - generate a sample by making a random change to the preceding sample
 - e.g. Markov Chain Monte Carlo (MCMC)
- These can be adapted for DBNs
 - Require several improvements to be practical
- E.g. Particle Filtering, an adaption of Likelihood Weighting

Work by generating random events based on the probabilities in the DBN and counting up the different



Particle Filtering

- First we generate a population of N samples from the prior distribution $\mathbf{P}(\mathbf{S}_0)$
- Then the update cycle is repeated for each time step:
 - sample, given the transition model $P(S_{t+1} | S_t)$
 - Each sample is weighted by the likelihood it assigns to the new observation, $\mathbf{P}(\mathbf{0}_{t+1} | \mathbf{s}_{t+1})$
 - ▶ The population is *resampled* to generate a new population of *N* samples
 - proportional to its weight. The new samples are unweighted.

inputs: **e**, the new incoming evidence *N*, the number of samples to be maintained local variables: W, a vector of weights of size N for i = 1 to N do $S[i] \leftarrow \text{sample from } \mathbf{P}(\mathbf{X}_1 | \mathbf{X}_0 = S[i])$ $W[i] \leftarrow \mathbf{P}(\mathbf{e} | \mathbf{X}_1 = S[i])$ $S \leftarrow \text{WEIGHTED-SAMPLE-WITH-REPLACEMENT}(N, S, W)$ return S

• Each sample is propagated forward by sampling the next state value s_{t+1} given the current value s_t for the

- Each new sample is selected from the current population; the probability that a particular sample is selected is

```
function PARTICLE-FILTERING(e, N, dbn) returns a set of samples for the next time step
           dbn, a DBN defined by \mathbf{P}(\mathbf{X}_0), \mathbf{P}(\mathbf{X}_1 | \mathbf{X}_0), and \mathbf{P}(\mathbf{E}_1 | \mathbf{X}_1)
persistent: S, a vector of samples of size N, initially generated from \mathbf{P}(\mathbf{X}_0)
                                                                 // step 1
                                                                 // step 2
                                                                                                   // step 3
```



Particle Filtering

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The particle filtering update cycle for the umbrella DBN with N = 10, showing the sample populations of each state. (a) At time t, 8 samples indicate rain and 2 indicate \neg rain. Each is propagated forward by sampling the next state through the transition model. At time t + 1, 6 samples indicate *rain* and 4 indicate \neg *rain*. (b) \neg *umbrella* is observed at *t* + 1. Each sample is weighted by its likelihood for the observation, as indicated by the size of the circles. (c) A new set of 10 samples is generated by weighted random selection from the current set, resulting in 2 samples that indicate *rain* and 8 that indicate \neg *rain*.





Particle Filtering

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- Then the update cycle is repeated for each time step:
 - Each sample is propagated forward by sampling the next state value s_{t+1} given the current value s_t for the sample, given the transition model $P(S_{t+1} | S_t)$
 - Each sample is weighted by the likelihood it assigns to the new observation, $P(o_{t+1} | s_{t+1})$
 - The population is *resampled* to generate a new population of N samples
 - Each new sample is selected from the current population; the probability that a particular sample is selected is proportional to its weight. The new samples are unweighted.
- Important property of Particle Filtering: it is *consistent* ▶ i.e. gives the correct probabilities as N tends to infinity $\lim N(\mathbf{s}_t | \mathbf{o}_{1,t}) / N = P(\mathbf{s}_t | \mathbf{o}_{1,t}) = \mathbf{f}_t$ $n \rightarrow \infty$

Also, efficient for many practical applications and is therefore widely used Maintains a good approximation to the true posterior using a constant number of samples Handles combinations of discrete and continuous variables Handles nonlinear and non-Gaussian models for continuous variables

where $N(\mathbf{s}_t | \mathbf{o}_{1:t})$ is the # of samples occupying state \mathbf{s}_t after observations $\mathbf{o}_{1:t}$ have been processed



Revisit HMM







_earning HMM Model



- Supervised: Learn transition and sensor models from labeled data
 - Access to both observations $\mathbf{0}_{1:t}$ and ground truth $\mathbf{s}_{0:t}$
 - Ground truth via manual labeling, addition sensor etc. Costly!
- Unsupervised: Learning from observations alone

 - These estimates can be used to update the models
 - The updated model provides new estimates, and the process iterates to convergence
- Training is also possible with prior expert-provided knowledge of some aspects of the model
- individually trained HMM can be combined to construct a larger HMM model e.g., of a complex activities with clear sub-activities structure

Russell, Stuart J., and Peter Norvig. "Artificial Intelligence: A Modern Approach"

Maximum likelihood estimate of the parameters of the HMM given the set of output sequences (intractable ⁽ⁱ⁾) Efficient practical approach: learning as a byproduct of inference - similar to Expectation Maximization algorithm - Inference provides an estimate of what transitions actually occurred and of what states generated the observations





A Common Algorithm: Baum–Welch

- It gives local optimum: gradient descent
- References:
 - https://www.youtube.com/watch?v=JRsdt05pMol
 - part of an excellent series on HMM
 - https://en.wikipedia.org/wiki/Baum-Welch algorithm
 - https://ocw.mit.edu/courses/aeronautics-andastronautics/16-410-principles-of-autonomy-anddecision-making-fall-2010/lecture-notes/ MIT16 410F10 lec21.pdf

Algorithm 1: The Baum-Welch algorithm Initialization: Θ_0 , $\{O_{1:T}\}$ Looping: for $l = 1, \ldots, l_{\text{max}}$ do 1. Forward-Backward calculations: $\alpha_1(i) = \pi_i b_i(O_1), \ \beta_T(i) = 1,$ $lpha_t(i) = \Bigl[\sum_{i=1}^K lpha_{t-1}(j) a_{ji} \Bigr] b_j(O_t), \; eta_t(i) = \sum_{i=1}^K a_{ij} b_j(O_{t+1}) eta_{t+1}(j)$ for $1 \leq i \leq K$, $1 \leq t \leq T-1$ 2. E-step: $\gamma_t(i) = \frac{\alpha_t(i)\beta_t(i)}{\sum_{j=1}^{K} \alpha_t(j)\beta_t(j)}, \ \xi_t(i,j) = \frac{\alpha_t(i)a_{ij}b_j(O_{t+1})\beta_{t+1}(j)}{\sum_{j=1}^{N} \sum_{j=1}^{N} \alpha_t(i)a_{ij}b_j(O_{t+1}\beta_t(j))}$ for $1 \le i \le K$, $1 \le j \le K$, $1 \le t \le T-1$ 3. M-step: $\pi_{i} = \frac{\gamma_{1}(i)}{\sum_{j=1}^{K} \gamma_{1}(j)}, \ a_{ij} = \frac{\sum_{t=1}^{T} \varepsilon_{t}(i,j)}{\sum_{k=1}^{K} \sum_{t=1}^{T} \varepsilon_{t}(i,k)}, \ w_{kd} = \frac{\sum_{t=1}^{T} \gamma_{t}(k,d)}{\sum_{t=1}^{T} \sum_{r=1}^{D} \gamma_{t}(k,r)}$ for $1 \le i \le K$, $1 \le j \le K$, $1 \le k \le K$, $1 \le d \le D$ end

Result: $\{\Theta_l\}_{l=0}^{l_{\max}}$



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Limitations of HMM

- Difficulty in representing multiple interacting activities (concurrent or interwoven)
- Incapable of capturing long-range or transitive dependencies of the observations due to its very strict independence assumptions (on the observations)
- Without significant training, an HMM may not be able to recognize all of the possible observation sequences that can be consistent with a particular activity.
- In practice:

 - Many activities are concurrent or interwoven Many activities may have non-deterministic natures - some steps of the activities may be done in any order



A More Flexible Alternative: Conditional Random Field (CRF)

- Addresses practical requirements that HMM cannot
- Two types of models
 - Generative: finds a joint probability distribution p(x, y) of a hidden variable y and an observed variables x
 - Discriminative: finds only the conditional probability p(y|x)
- Both can be used to find a hidden state transition from observation sequences
- What type is HMM?



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- Both can be used to find a hidden state transition from observation sequences
- What type is HMM? Generative!
- CRF: a conditional distribution with an associated graphical structure • Attempts to find only the conditional probability p(y | x), i.e. discriminative approach Allows for arbitrary, non-independent relationships among the observation sequences
- - Relaxes the independence assumptions
 - the hidden state probabilities may depend on the past and even future observations Modeled as an undirected acyclic graph with two types of nodes: observed & unobserved





https://towardsdatascience.com/conditional-random-fields-explained-e5b8256da776



- Represented by a graph G = (V, E)
 - nodes represent random variables
 - edges collectively represent the dependencies between them
- The graph can be factorized into J different cliques or factors
 - ▶ each governed by a factor function Φ_j with its scope being a subset of random variables D_i where $\Phi_i(d_i) > 0 \forall d_i \in D_i$
- The unnormalized joint probability of the variables is the product of all the factor function, so that

$$Pr(A = a, B = b, C = c, D = d) = \frac{\phi_1(a, b)\phi_2}{\sum_{a'b'c'd'}} \phi_1(a', b')$$







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- Gibb's Notation: represent the joint as a Gibbs distribution by operating on factor functions in log space by using β(d ;)= log(φ(d ;))
 - X is the set of all the random variables in the graph



Gibbs,
$$P(x) = \frac{e^{-E(x)}}{Z}$$
, where $Z = \sum_{x' \in X} e^{-E(x)}$







Conditional Random Field Model

- CRF is special case of this MRF wherein the graph satisfies the property:

where Y₁~Y₂ signifies that Y₁ and Y₂ are neighbors in the graph."

- A variable's neighboring nodes or variables are also called the <u>Markov Blanket</u> of that variable
- One such graph that satisfies the above property is the chain-structured graph shared below:



Assume a MRF divides into two sets of random variables Y (unobserved) and X (observed) respectively

"When we condition the graph on X globally i.e. when the values of random variables in X is fixed or given, all the random variables in set Y follow the Markov property $p(Y_{11}|X,Y_{2}, u \neq v) = p(Y_{11}|X,Y_{2}, Y_{11} \sim Y_{2})$,



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Making Inferences with CRF

- CRF is a discriminative model i.e. it models the conditional probability P(Y|X)▶ i.e. X is always given or observed
- Therefore the graph ultimately reduces to a simple chain



- called the evidence (sensor observations) and label variables respectively
- Training problem: maximizing the log likelihood wrt all model parameters
 - See "Conditional Random Fields Explained" by Aditya Prasad @ the URL in left bottom footer
 - Many off-the-shelf packages

As we condition upon X and we are trying to find the corresponding Y_i for every X_i, X and Y are also



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CRF Version of the Eating Activity Example



Kim, Eunju, Sumi Helal, and Diane Cook. "Human activity recognition and pattern discovery."



Practical Resources

- Pomegranate package for Python
 - Many probabilistic models: HMM, Naive Bayes, and various others...
 - Learning Research 18, no. 1 (2017): 5992-5997.
 - https://jmlr.org/papers/volume18/17-636/17-636.pdf
 - https://github.com/jmschrei/pomegranate and https://pomegranate.readthedocs.io
- **pgmpy** python library for working with Probabilistic Graphical Models.
 - Many probabilistic graphs models: Dynamic Bayesian Networks, Naive Bayes etc.
 - https://github.com/pgmpy/pgmpy and http://pgmpy.org
- Fun with Hidden Markov Models in PyTorch (by Loren Lugosch)
 - recognition." Proceedings of the IEEE 77, no. 2 (1989): 257-286.
 - https://www.cs.cmu.edu/~cga/behavior/rabiner1.pdf
 - https://colab.research.google.com/drive/1IUe9lfoliQsL49atSOgxnCmMR_zJazKI
- python-crfsuite python binding for CRFSuite
 - An implementation of Conditional Random Fields (CRFs) for labeling sequential data
 - https://github.com/chokkan/crfsuite

scikit-learn (https://scikit-learn.org) provides many general classification and regression algorithms

Schreiber, Jacob. "Pomegranate: fast and flexible probabilistic modeling in python." The Journal of Machine

▶ Based on: Rabiner, Lawrence R. "A tutorial on hidden Markov models and selected applications in speech



Practical Resources (contd.)

- **FilterPy** package for Python
 - Python library that implements a number of Bayesian filters, most notably Kalman filters
 - https://github.com/rlabbe/filterpy
 - Excellent related interactive book: Kalman and Bayesian Filter in Python
 - https://github.com/rlabbe/Kalman-and-Bayesian-Filters-in-Python/
- **pykalman** library for Python
 - Kalman Filter, Kalman Smoother, and EM
 - https://pykalman.github.io and https://github.com/pykalman/pykalman
- TinyEKF Lightweight C/C++ Extended Kalman Filter with Python https://github.com/simondlevy/TinyEKF
- **tsBNgen** for generating synthetic time series data in Python (from Prof. Pottie's group)
 - Python package to generate time series data based on an arbitrary Bayesian Network structures
 - Dynamic Bayesian Network Structure." arXiv preprint arXiv:2009.04595 (2020).
 - <u>https://arxiv.org/pdf/2009.04595.pdf</u>
 - https://github.com/manitadayon/tsBNgen#Features

Tadayon, Manie, and Greg Pottie. "tsBNgen: A Python Library to Generate Time Series Data from an Arbitrary

